Haplo Stats

(version 1.5.0)

Statistical Methods for Haplotypes When Linkage Phase is Ambiguous

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1 Introduction

Haplo Stats is a suite of R routines for the analysis of indirectly measured haplotypes. The statistical methods assume that all subjects are unrelated and that haplotypes are ambiguous (due to unknown linkage phase of the genetic markers), while also allowing for missing alleles.

The user-level functions in Haplo Stats are:

- *haplo.em:* for the estimation of haplotype frequencies and posterior probabilities of haplotype pairs for each subject, conditional on the observed marker data
- *haplo.glm:* generalized linear models for the regression of a trait on haplotypes, with the option of including covariates and interactions
- *haplo.score:* score statistics to test associations between haplotypes and a variety of traits, including binary, ordinal, quantitative, and Poisson
- haplo.score.slide: haplo.score computed on sub-haplotypes of a larger region
- sequentially scan markers in enlarging a haplotype for association with a trait
- *haplo.cc:* run a combined analysis for haplotype frequencies, scores, and regression results for a case-control study
- haplo.power.qt/haplo.power.cc: power or sample size calculatins for quantitative or binary trait
- *haplo.scan:* search for a trait locus for all sizes of sub-haplotypes within a fixed maximum window width for all markers in a region
- haplo.design: create a design matrix for haplotype effects

This manual explains the basic and advanced usage of these routines, with guidelines for running the analyses and interpreting results. We provide many of these details in the function help pages, which are accessed within an R session using help(haplo.em), for example. We also provide brief examples in the help files, which can be run in the R session with example(haplo.em).

1.1 Updates

This version of Haplo Stats contains updates to *haplo.glm* in section 6 and new methods written for it that resemble glm class methods. These methods include residuals, fitted.values, vcov, and anova, and they are detailed in section 7. For full history of updates see the NEWS file, or type *news(package="haplo.stats")* in the R command prompt.

1.2 Operating System and Installation

Haplo Stats version 1.5.0 is written for R (version 2.14.0). It has been uploaded to the Comprehensive R Archive Network (CRAN), and is made available on various operating systems through CRAN. Package installation within R is made simple from within R using *install.packages("haplo.stats")*, but other procedures for installing R packages can be found at the R project website (http://www.r-project.org).

1.3 R Basics

For users who are new to the R environment, we demonstrate some basic concepts. In the following example we create a vector of character alleles and use the *table* function to get allele counts. We first show how to save the results of table(snp) into an R session variable, tab. We show that tab is an object of the *table* class, and use the print and summary methods that are defined for table objects. Note that when you enter just tab or table(snp) at the prompt, the print method is invoked.

```
> snp <- c("A", "T", "T", "A", "A", "T", "T")
> tab <- table(snp)</pre>
> tab
snp
A T
34
> class(tab)
[1] "table"
> print.table(tab)
snp
ΑT
34
> summary(tab)
Number of cases in table: 7
Number of factors: 1
> tab[2]
Т
4
> table(snp)
snp
A T
34
```

The routines in *haplo.stats* are computationally intensive and return lots of information in the returned object. Therefore, we assign classes to the returned objects and provide various methods for each of them.

2 Data Setup

We first show some typical steps when you first load a package and look for details on a function of interest. In the sample code below, we load *haplo.stats*, check which functions are available in the package, view a help file, and run the example that is within the help file.

```
> # load the library, load and preview at demo dataset
```

```
> library(haplo.stats)
```

> ls(name="package:haplo.stats")

```
> help(haplo.em)
```

```
> example(haplo.em)
```

2.1 Example Data

The haplo.stats package contains three example data sets. The primary data set used in this manual is named (hla.demo), which contains 11 loci from the HLA region on chromosome 6, with covariates, qualitative, and quantitative responses. Within /haplo.stats/data/hla.demo.tab the data is stored in tab-delimited format. Typically data stored in this format can be read in using read.table(). Since the data is provided in the package, we load the data in R using data() and view the names of the columns. Then to make the columns of hla.demo accessible without typing it each time, we attach it to the current session.

```
> # load and preview demo dataset stored in ~/haplo.stats/data/hla.demo.tab
> data(hla.demo)
> names(hla.demo)
 [1] "resp"
                 "resp.cat"
                            "male"
                                        "age"
                                                    "DPB.a1"
                                                                "DPB.a2"
 [7] "DPA.a1"
                 "DPA.a2"
                                        "DMA.a2"
                                                    "DMB.a1"
                                                                "DMB.a2"
                            "DMA.a1"
                 "TAP1.a2"
[13] "TAP1.a1"
                            "TAP2.a1"
                                        "TAP2.a2"
                                                    "DQB.a1"
                                                               "DQB.a2"
                                                                "B.a2"
[19] "DQA.a1"
                 "DQA.a2"
                            "DRB.a1"
                                        "DRB.a2"
                                                    "B.a1"
[25] "A.a1"
                 "A.a2"
> # attach hla.demo to make columns available in the session
```

```
> attach(hla.demo)
```

The column names of *hla.demo* are shown above. They are defined as follows:

- resp: quantitative antibody response to measles vaccination
- resp.cat: a factor with levels "low", "normal", "high", for categorical antibody response
- male: gender code with 1="male", 0="female"
- age: age (in months) at immunization

The remaining columns are genotypes for 11 HLA loci, with a prefix name (e.g., "DQB") and a suffix for each of two alleles (".a1" and ".a2"). The variables in *hla.demo* can be accessed by typing *hla.demo*\$ before their names, such as *hla.demo*\$resp. Alternatively, it is easier for these examples to attach *hla.demo*, (as shown above with *attach()*) so the variables can be accessed by simply typing their names.

2.2 Creating a Genotype Matrix

Many of the functions require a matrix of genotypes, denoted below as geno. This matrix is arranged such that each locus has a pair of adjacent columns of alleles, and the order of columns corresponds to the order of loci on a chromosome. If there are K loci, then the number of columns of geno is 2K. Rows represent the alleles for each subject. For example, if there are three loci, in the order A-B-C, then the 6 columns of geno would be arranged as A.a1, A.a2, B.a1, B.a2, C.a1, C.a2. For illustration, three of the loci in *hla.demo* will be used to demonstrate some of the functions. Create a separate data frame for 3 of the loci, and call this geno. Then create a vector of labels for the loci.

> geno <- hla.demo[,c(17,18,21:24)]

> label <-c("DQB","DRB","B")</pre>

The *hla.demo* data already had alleles in two columns for each locus. For many SNP datasets, the data is in a one column format, giving the count of the minor allele. To assist in converting this format to two columns, a function named *geno1to2* has been added to the package. See its help file for more details.

2.3 Preview Missing Data: summaryGeno

Before performing a haplotype analysis, the user will want to assess missing genotype data to determine the completeness of the data. If many genotypes are missing, the functions may take a long time to compute results, or even run out of memory. For these reasons, the user may want to remove some of the subjects with a lot of missing data. This step can be guided by using the *summaryGeno* function, which checks for missing allele information and counts the number of potential haplotype pairs that are consistent with the observed data (see the Appendix for a description of this counting scheme).

The codes for missing alleles are defined by the parameter *miss.val*, a vector to define all possible missing value codes. Below, the result is saved in *geno.desc*, which is a data frame, so individual rows may be printed. Here we show the results for subjects 1-10, 80-85, and 135-140, some of which have missing alleles.

```
> geno.desc <- summaryGeno(geno, miss.val=c(0,NA))
> print(geno.desc[c(1:10,80:85,135:140),])
```

	missing0	missing1	missing2	N.enum.rows
1	3	0	0	4
2	3	0	0	4
3	3	0	0	4
4	3	0	0	2
5	3	0	0	4
6	3	0	0	2
7	3	0	0	4
8	3	0	0	2
9	3	0	0	2
10	3	0	0	1

80	3	0	0	4
81	2	0	1	1800
82	3	0	0	2
83	3	0	0	1
84	3	0	0	2
85	3	0	0	4
135	3	0	0	4
136	3	0	0	2
137	1	0	2	129600
138	3	0	0	4
139	3	0	0	4
140	3	0	0	4

The columns with 'loc miss-' illustrate the number of loci missing either 0, 1, or 2 alleles, and the last column, num_enum_rows, illustrates the number of haplotype pairs that are consistent with the observed data. In the example above, subjects indexed by rows 81 and 137 have missing alleles. Subject #81 has one locus missing two alleles, while subject #137 has two loci missing two alleles. As indicated by num_enum_rows, subject #81 has 1,800 potential haplotype pairs, while subject #137 has nearly 130,000.

The 130,000 haplotype pairs is considered a large number, but haplo.em, haplo.score, and haplo.glm complete in roughly 3-6 minutes (depending on system limits or control parameter settings). If person #137 were removed, the methods would take less than half that time. It is preferred to keep people if they provide information to the analysis, given that run time and memory usage are not too much of a burden.

When a person has no genotype information, they do not provide information to any of the methods in *haplo.stats*. Furthermore, they cause a much longer run time. Below, using the *table* function on the third column of *geno.desc*, we can tabulate how many people are missing two alleles at any at of the three loci. If there were people missing two alleles at all three loci, they should be removed. The second command below shows how to make an index of which people to remove from *hla.demo* because they are missing all their alleles.

```
> # find if there are any people missing all alleles
> table(geno.desc[,3])
0 1 2
218 1 1
> ## create an index of people missing all alleles
> miss.all <- which(geno.desc[,3]==3)
> # use index to subset hla.demo
> hla.demo.updated <- hla.demo[-miss.all,]</pre>
```

2.4 Random Numbers and Setting Seed

Random numbers are used in several of the functions (e.g., to determine random starting frequencies within *haplo.em*, and to compute permutation p-values in *haplo.score*). To reproduce calculations involving random numbers, we set the seed values stored in a vector called *.Random.seed*. This

vector can be set using *set.seed()* before any function which uses random numbers. Section 6.7.2 shows one example of setting the seed for *haplo.glm*. We illustrate setting the seed below.

```
> # this is how to set the seed for reproducing results where haplo.em is
> # involved, and also if simulations are run. In practice, don't reset seed.
> seed <- c(17, 53, 1, 40, 37, 0, 62, 56, 5, 52, 12, 1)
> set.seed(seed)
```

3 Haplotype Frequency Estimation: haplo.em

3.1 Algorithm

For genetic markers measured on unrelated subjects, with linkage phase unknown, haplo.em computes maximum likelihood estimates of haplotype probabilities. Because there may be more than one pair of haplotypes that are consistent with the observed marker phenotypes, posterior probabilities of haplotype pairs for each subject are also computed. Unlike the usual EM which attempts to enumerate all possible haplotype pairs before iterating over the EM steps, our *progressive insertion* algorithm progressively inserts batches of loci into haplotypes of growing lengths, runs the EM steps, trims off pairs of haplotypes per subject when the posterior probability of the pair is below a specified threshold, and then continues these insertion, EM, and trimming steps until all loci are inserted into the haplotype. The user can choose the batch size. If the batch size is chosen to be all loci, and the threshold for trimming is set to 0, then this reduces to the usual EM algorithm. The basis of this progressive insertion algorithm is from the "snphap" software by David Clayton[1]. Although some of the features and control parameters of *haplo.em* are modeled after *snphap*, there are substantial differences, such as extension to allow for more than two alleles per locus, and some other nuances on how the algorithm is implemented.

3.2 Example Usage

We use haplo.em on geno for the 3 loci defined above and save the result in an object named save.em, which has the haplo.em class. The print method would normally print all 178 haplotypes from save.em, but to keep the results short for this manual, we give a quick glance of the output by using the option nlines=10, which prints only the first 10 haplotypes of the full results. The nlines parameter has been employed in some of Haplo Stats' print methods for when there are many haplotypes. In practice, it is best to exclude this parameter so that the default will print all results.

```
> save.em <- haplo.em(geno=geno, locus.label=label, miss.val=c(0,NA))
> names(save.em)
```

[1]	"lnlike"		"lr"	"df.lr"
[5]	"hap.prob"	"hap.prob.noLD"	"converge"	"locus.label"
[9]	"indx.subj"	"subj.id"	"post"	"hap1code"
[13]	"hap2code"	"haplotype"	"nreps"	"rows.rem"
[17]	"max.pairs"	"control"		

```
> print(save.em, nlines=10)
```

==														
	Haplotypes													
==														
	DQB	DRB	В	hap.freq										
1	21	1	8	0.00232										
2	21	2	7	0.00227										
3	21	2	18	0.00227										
4	21	3	8	0.10408										
5	21	3	18	0.00229										
6	21	3	35	0.00570										
7	21	3	44	0.00378										
8	21	3	45	0.00227										
9	21	3	49	0.00227										
10	21	3	57	0.00227										
==	====:	====	===:											
					Details									
==	===== , .,	====:	====											
				47.675										
lr	lr stat for no LD = 632.5085 , df = 125 , p-val = 0													

Explanation of Results

The print methods shows the haplotypes and their estimated frequencies, followed by the final log-likelihood statistic and the lr stat for no LD, which is the likelihood ratio test statistic contrasting the *lnlike* for the estimated haplotype frequencies versus the *lnlike* under the null assuming that alleles from all loci are in linkage equilibrium. We note that the trimming by the progressive insertion algorithm can invalidate the lr stat and the degrees of freedom (*df*).

3.3 Summary Method

The summary method for a haplo.em object on save.em shows the list of haplotypes per subject, and their posterior probabilities:

```
> summary(save.em, nlines=7)
```

=:		Subier		type Codes	===== 2nd	Posterior	Probabilities	
=		========	=======================================	======================================	=====	=======================================	=================	
	subj.id	hap1code	hap2code	posterior				
1	1	58	78	1.00000				
2	2	13	143	0.12532				
3	2	17	138	0.87468				
4	3	25	168	1.00000				
5	4	13	39	0.28621				
6	4	17	38	0.71379				
7	5	55	94	1.00000				

Number of haplotype pairs: max vs used										
	=====			====						
х	1	2	3	84	135					
1	18	0	0	0	0					
2	50	4	0	0	0					
4	116	29	1	0	0					
1800	0	0	0	1	0					
129600	0	0	0	0	1					

Explanation of Results

The first part of the summary output lists the subject id (row number of input geno matrix), the codes for the haplotypes of each pair, and the posterior probabilities of the haplotype pairs. The second part gives a table of the maximum number of pairs of haplotypes per subject, versus the number of pairs used in the final posterior probabilities. The haplotype codes remove the clutter of illustrating all the alleles of the haplotypes, but may not be as informative as the actual haplotypes themselves. To see the actual haplotypes, use the show.haplo=TRUE option, as in the following example.

> # show full haplotypes, instead of codes

> summary(save.em, show.haplo=TRUE, nlines=7)

Subjects: Haplotype Codes and Posterior Probabilities													
	subj.id	hap1	.DQB	hap1	L.DRB	hap1.B	hap2.DQ	 B hap2	.DRB	hap2.B	posterior		
58	1		31		11	61	3	2	4	62	1.00000		
13	2		21		7	7	6	2	2	44	0.12532		
17	2		21		7	44	6	2	2	7	0.87468		
25	3		31		1	27	6	3	13	62	1.00000		
13.1	4		21		7	7	3	1	7	44	0.28621		
17.1	4		21		7	44	3	1	7	7	0.71379		
55	5		31		11	51	4	2	8	55	1.00000		
Number of haplotype pairs: max vs used													
x	1	2	3	84 1	L35								
1	18	0	0	0	0								
2	50	4	0	0	0								
4	116	29	1	0	0								
180	0 0	0	0	1	0								
129	600 0	0	0	0	1								

3.4 Control Parameters for haplo.em

A set of control parameters can be passed together to haplo.em as the "control" argument. This is a list of parameters that control the EM algorithm based on progressive insertion of loci. The default values are set by a function called haplo.em.control (see help(haplo.em.control) for a complete description). Although the user can accept the default values, there are times when control parameters may need to be adjusted.

These parameters are defined below:

- insert.batch.size: Number of loci to be inserted in a single batch.
- **min.posterior:** Minimum posterior probability of haplotype pair, conditional on observed marker genotypes. Posteriors below this minimum value will have their pair of haplotypes "trimmed" off the list of possible pairs.
- **max.iter:** Maximum number of iterations allowed for the EM algorithm before it stops and prints an error.
- **n.try:** Number of times to try to maximize the *lnlike* by the EM algorithm. The first try will use, as initial starting values for the posteriors, either equal values or uniform random variables, as determined by *random.start*. All subsequent tries will use uniform random values as initial starting values for the posterior probabilities.
- max.haps.limit: Maximum number of haplotypes for the input genotypes. Within haplo.em, the first step is to try to allocate the sum of the result of geno.count.pairs(), if that exceeds max.haps.limit, start by allocating max.haps.limit. If that is exceeded in the progressive-insertions steps, the C function doubles the memory until it can longer request more.

One reason to adjust control parameters is for finding the global maximum of the log-likelihood. It can be difficult in particular for small sample sizes and many possible haplotypes. Different maximizations of the log-likelihood may result in different results from haplo.em, haplo.score, or haplo.glm when rerunning the analyses. The algorithm uses multiple attempts to maximize the log-likelihood, starting each attempt with random starting values. To increase the chance of finding the global maximum of the log-likelihood, the user can increase the number of attempts (n.try), increase the batch size (insert.batch.size), or decrease the trimming threshold for posterior probabilities (min.posterior).

Another reason to adjust control parameters is when the algorithm runs out of memory because there are too many haplotypes. If max.haps.limit is exceeded when a batch of markers is added, the algorithm requests twice as much memory until it runs out. One option is to set max.haps.limit to a different value, either to make haplo.em request more memory initially, or to request more memory in smaller chunks. Another solution is to make the algorithm trim the number of haplotypes more aggressively by decreasing insert.batch.size or increasing min.posterior. Any changes to these parameters should be made with caution, and not drastically different from the default values. For instance, the default for min.posterior used to be 1e-7, and in some rare circumstances with many markers in only moderate linkage disequilibrium, some subjects had all their possible haplotype pairs trimmed. The default is now set at 1e-9, and we recommend not increasing min.posterior much greater than 1e-7.

The example below gives the command for increasing the number of tries to 20, and the batch size to 2, since not much more can be done for three markers.

```
> # demonstrate only the syntax of control parameters
> save.em.try20 <- haplo.em(geno=geno, locus.label=label, miss.val=c(0, NA),
+ control = haplo.em.control(n.try = 20, insert.batch.size=2))
```

3.5 Haplotype Frequencies by Group Subsets

To compute the haplotype frequencies for each level of a grouping variable, use the function *haplo.group*. The following example illustrates the use of a binomial response based on *resp.cat*, *y.bin*, that splits the subjects into two groups.

```
> ## run haplo.em on sub-groups
> ## create ordinal and binary variables
> y.bin <- 1*(resp.cat=="low")</pre>
> group.bin <- haplo.group(y.bin, geno, locus.label=label, miss.val=0)</pre>
> print(group.bin, nlines=15)
                           _____
                   Counts per Grouping Variable Value
      _____
group
 0
    1
157 63
                    Haplotype Frequencies By Group
_____
  DQB DRB B
             Total y.bin.0 y.bin.1
   21
         8 0.00232 0.00335
       1
1
                             NΑ
   21 10 8 0.00181 0.00318
2
                             NA
3
   21 13 8 0.00274
                             NA
                      ΝA
4
   21
       2 18 0.00227 0.00318
                             NΑ
5
   21 2 7 0.00227 0.00318
                             NA
6
   21
       3 18 0.00229 0.00637
                             NA
7
   21 3 35 0.00570 0.00639
                             NA
8
   21
       3 44 0.00378 0.00333 0.01587
9
   21
       3 45 0.00227
                      NA
                             NΑ
10
   21 3 49 0.00227
                      NA
                             NA
11
   21
       3 57 0.00227
                      NA
                             NA
12 21
       3 70 0.00227
                      NA 0.00000
       3 8 0.10408 0.06974 0.19048
13
   21
14
   21
       4 62 0.00455 0.00637
                             NΑ
15
   21
       7 13 0.01072
                      NA 0.02381
```

Explanation of Results

The group.bin object can be very large, depending on the number of possible haplotypes, so only

a portion of the output is illustrated above (limited again by *nlines*). The first section gives a short summary of how many subjects appear in each of the groups. The second section is a table with the following columns:

- The first column gives row numbers.
- The next columns (3 in this example) illustrate the alleles of the haplotypes.
- Total are the estimated haplotype frequencies for the entire data set.
- The last columns are the estimated haplotype frequencies for the subjects in the levels of the group variable (*y.bin=0* and *y.bin=1* in this example). Note that some haplotype frequencies have an NA, which appears when the haplotypes do not occur in the subgroups.

4 Power and Sample Size for Haplotype Association Studies

It is known that using haplotypes has greater power than single-markers to detect genetic association in some circumstances. There is little guidance, however, in determining sample size and power under different circumstances, some of which include: marker type, dominance, and effect size. The *haplo.stats* package now includes functions to calculate sample size and power for haplotype association studies, which is flexible to handle these multiple circumstances.

Based on work in Schaid 2005[2], we can take a set of haplotypes with their population frequencies, assign a risk to a subset of the haplotypes, then determine either the sample size to achieve a stated power, or the power for a stated sample size. Sample size and power can be calculated for either quantitative traits or case-control studies.

4.1 Quantitative Traits: haplo.power.qt

We assume that quantitative traits will be modeled by a linear regression. Some well-known tests for association between haplotypes and the trait include score statistics[3] and an F-test[4]. For both types of tests, power depends on the amount of variance in the trait that is explained by haplotypes, or a multiple correlation coefficient, R^2 . Rather than specifying the haplotype coefficients directly, we calculate the vector of coefficients based on an R^2 value.

In the example below, we load an example set of haplotypes that contain 5 markers, and specify the indices of the at-risk haplotypes; in this case, whichever haplotype has allele 1 at the 2nd and 3rd markers. We set the first haplotype (most common) as the baseline. With these values we calculate the vector of coefficients for haplotype effects from find.haplo.beta.qt using an $R^2 = 0.01$. Next, we use haplo.power.qt to calculate the sample size for the set of haplotypes and their coefficients, type-I error (alpha) set to 0.05, power at 80%, and the same mean and variance used to get haplotype coefficients. Then we use the sample size needed for 80% power for un-phased haplotypes (2,826) to get the power for both phased and un-phased haplotypes.

```
> # load a set of haplotypes (hap-1 from Schaid 2005)
>
> data(hapPower.demo)
> ##### an example using save.em hla markers may go like this.
> # keep <- which(save.em$hap.prob > .004) # get an index of non-rare haps
```

```
> # hfreq <- save.em$hap.prob[keep]</pre>
> # hmat <- save.em$haplotype[keep,]</pre>
> # hrisk <- which(hmat[,1]==31 & hmat[,2]==11) # contains 3 haps with freq=.01
> # hbase <- 4 # 4th hap has mas freq of .103
> ####
>
> ## separate the haplotype matrix and the frequencies
> hmat <- hapPower.demo[,-6]</pre>
> hfreq <- hapPower.demo[,6]</pre>
> ## Define risk haplotypes as those with "1" allele at loc2 and loc3
> hrisk <- which(hmat$loc.2==1 & hmat$loc.3==1)</pre>
> # define index for baseline haplotype
> hbase <- 1
> hbeta.list <- find.haplo.beta.qt(haplo=hmat, haplo.freq=hfreq, base.index=hbase,</pre>
                                 haplo.risk=hrisk, r2=.01, y.mu=0, y.var=1)
> hbeta.list
$r2
[1] 0.01
$beta
 [7] 0.0000000 0.27636731 0.0000000 0.27636731 0.0000000 0.0000000
[13] 0.00000000 0.00000000 0.00000000 0.27636731 0.27636731 0.00000000
[19] 0.0000000 0.0000000 0.0000000
$base.index
[1] 1
$haplo.risk
[1] 8 10 16 17
> ss.qt <- haplo.power.qt(hmat, hfreq, hbase, hbeta.list$beta,</pre>
+
                         y.mu=0, y.var=1, alpha=.05, power=.80)
> ss.qt
$ss.phased.haplo
[1] 2091
$ss.unphased.haplo
[1] 2826
$power.phased.haplo
[1] 0.8
$power.unphased.haplo
[1] 0.8
```

```
> power.qt <- haplo.power.qt(hmat, hfreq, hbase, hbeta.list$beta,
+ y.mu=0, y.var=1, alpha=.05, sample.size=2826)
> power.qt
$ss.phased.haplo
[1] 2826
$ss.unphased.haplo
[1] 2826
$power.phased.haplo
[1] 0.9282451
$power.unphased.haplo
[1] 0.8000592
```

4.2 Case-Control Studies: haplo.power.cc

The steps to compute sample size and power for case-control studies is similar to the steps for quantitative traits. If we assume a log-additive model for haplotype effects, the haplotype coefficients can be specified first as odds ratios (OR), and then converted to logistic regression coefficients according to log(OR).

In the example below, we assume the same baseline and risk haplotypes defined in section 4.1, give the risk haplotypes an odds ratio of 1.50, and specify a population disease prevalance of 10%. We also assume cases make up 50% (*case.frac*) of the study's subjects. We first compute the sample size for this scenario for Type-I error (alpha) at 0.05 and 80% power, and then compute power for the sample size required for un-phased haplotypes (4, 566).

```
> ## get power and sample size for quantitative response
> ## get beta vector based on odds ratios
>
> cc.OR <- 1.5
> # determine beta regression coefficients for risk haplotypes
>
> hbeta.cc <- numeric(length(hfreq))</pre>
> hbeta.cc[hrisk] <- log(cc.OR)</pre>
> # Compute sample size for stated power
>
> ss.cc <- haplo.power.cc(hmat, hfreq, hbase, hbeta.cc, case.frac=.5, prevalence=.1,
                 alpha=.05, power=.8)
+
> ss.cc
$ss.phased.haplo
[1] 3454
$ss.unphased.haplo
```

```
[1] 4566
$power.phased.haplo
[1] 0.8
$power.unphased.haplo
[1] 0.8
> # Compute power for given sample size
> power.cc <- haplo.power.cc(hmat, hfreq, hbase, hbeta.cc, case.frac=.5, prevalence=.1,
                 alpha=.05, sample.size=4566)
+
> power.cc
$ss.phased.haplo
[1] 4566
$ss.unphased.haplo
[1] 4566
$power.phased.haplo
[1] 0.9206568
$power.unphased.haplo
[1] 0.8000695
```

5 Haplotype Score Tests: haplo.score

The *haplo.score* routine is used to compute score statistics to test association between haplotypes and a wide variety of traits, including binary, ordinal, quantitative, and Poisson. This function provides several different global and haplotype-specific tests for association and allows for adjustment for non-genetic covariates. Haplotype effects can be specified as additive, dominant, or recessive. This method also has an option to compute permutation p-values, which may be needed for sparse data when distribution assumptions may not be met. Details on the background and theory of the score statistics can be found in Schaid et al.[3].

5.1 Quantitative Trait Analysis

First, we assess a haplotype association with a quantitative trait in hla.demo called resp. To tell haplo.score the trait is quantitative, specify the parameter trait.type="gaussian" (a reminder that a gaussian distribution is assumed for the error terms). The other arguments, all set to default values, are explained in the help file. Note that rare haplotypes can result in unstable variance estimates, and hence unreliable test statistics for rare haplotypes. We restrict the analysis to get scores for haplotypes with a minimum sample count using min.count=5. For more explanation on handling rare haplotypes, see section 5.6. Below is an example of running haplo.score with a

quantitative trait, then viewing the results using the *print* method for the *haplo.score* class. (again, output shortened by *nlines*).

> # score statistics w/ Gaussian trait > score.gaus.add <- haplo.score(resp, geno, trait.type="gaussian",</pre> + min.count=5, locus.label=label, simulate=FALSE) > print(score.gaus.add, nlines=10) _____ Haplotype Effect Model: additive _____ Global Score Statistics _____ global-stat = 30.6353, df = 18, p-val = 0.03171 _____ Haplotype-specific Scores _____ DQB DRB B Hap-Freq Hap-Score p-val [1,] 21 8 0.10408 -2.39631 0.01656 3 [2,] 31 4 44 0.02849 -2.24273 0.02491 [3,] 51 1 44 0.01731 -0.99357 0.32043 [4,] 63 13 44 0.01606 -0.84453 0.39837 [5,] 63 2 7 0.01333 -0.50736 0.6119 [6,] 32 4 60 0.0306 -0.46606 0.64118 [7,] 21 7 44 0.02332 -0.41942 0.67491 [8,] 62 2 44 0.01367 -0.26221 0.79316 [9,] 62 2 18 0.01545 -0.21493 0.82982 [10,] 51 1 27 0.01505 0.01539 0.98772

Explanation of Results

First, the model effect chosen by haplo.effect is printed across the top. The section Global Score Statistics shows results for testing an overall association between haplotypes and the response. The global-stat has an asymptotic χ^2 distribution, with degrees of freedom (df) and p-value as indicated. Next, Haplotype-specific scores are given in a table format. The column descriptions are as follows:

- The first column gives row numbers.
- The next columns (3 in this example) illustrate the alleles of the haplotypes.
- Hap-Freq is the estimated frequency of the haplotype in the pool of all subjects.
- *Hap-Score* is the score for the haplotype, the results are sorted by this value. Note, the score statistic should not be interpreted as a measure of the haplotype effect.

• *p*-val is the asymptotic χ_1^2 p-value, calculated from the square of the score statistic.

5.2 Binary Trait Analysis

Let us assume that "low" responders are of primary interest, so we create a binary trait that has values of 1 when resp.cat is "low", and 0 otherwise. Then in haplo.score specify the parameter trait.type="binomial".

```
> # scores, binary trait
> y.bin <- 1*(resp.cat=="low")</pre>
> score.bin <- haplo.score(y.bin, geno, trait.type="binomial",</pre>
                    x.adj = NA, min.count=5,
+
+
                    haplo.effect="additive", locus.label=label,
                    miss.val=0, simulate=FALSE)
+
> print(score.bin, nlines=10)
_____
                  Haplotype Effect Model: additive
                                _____
     _____
                     Global Score Statistics
_____
                   _____
global-stat = 33.70125, df = 18, p-val = 0.01371
                               -----
                    Haplotype-specific Scores
                _____
                                         -------
    DQB DRB B Hap-Freq Hap-Score p-val
[1,] 62
       2
          7 0.05098 -2.19387 0.02824
[2,] 51
          35 0.03018 -1.58421 0.11315
       1
[3,] 63
       13 7 0.01655 -1.56008 0.11874
[4,] 21
       7
          7 0.01246 -1.47495 0.14023
[5,] 32
          7 0.01678 -1.00091 0.31687
       4
[6,] 32
       4
          62 0.02349 -0.6799
                           0.49657
[7,] 51
          27 0.01505 -0.66509 0.50599
       1
[8,] 31
       11 35 0.01754 -0.5838
                           0.55936
[9,] 31
       11 51 0.01137 -0.43721 0.66196
[10,] 51
          44 0.01731 0.00826
                           0.99341
       1
```

5.3 Ordinal Trait Analysis

To create an ordinal trait, here we convert *resp.cat* (described above) to numeric values, *y.ord* (with levels 1, 2, 3). For *haplo.score*, use *y.ord* as the response variable, and set the parameter *trait.type* = "ordinal".

```
> # scores w/ ordinal trait
> y.ord <- as.numeric(resp.cat)</pre>
> score.ord <- haplo.score(y.ord, geno, trait.type="ordinal",</pre>
                 x.adj = NA, min.count=5,
+
+
                 locus.label=label,
                 miss.val=0, simulate=FALSE)
+
> print(score.ord, nlines=7)
                  Haplotype Effect Model: additive
_____
 _____
                     Global Score Statistics
    _____
                       _____
global-stat = 15.23209, df = 18, p-val = 0.64597
                    Haplotype-specific Scores
                                        _____
   DQB DRB B Hap-Freq Hap-Score p-val
[1,] 32 4
         62 0.02349 -2.17133 0.02991
[2,] 21 3
         8 0.10408 -1.34661 0.17811
[3,] 32 4
         7 0.01678 -1.09487 0.27357
[4,] 62 2
         7 0.05098 -0.96874 0.33268
[5,] 21
      7
         44 0.02332 -0.83747 0.40233
[6,] 63
        7 0.01655
                  -0.80787 0.41917
      13
[7,] 21
      7
         7 0.01246 -0.63316 0.52663
```

Warning for Ordinal Traits

When analyzing an ordinal trait with adjustment for covariates (using the x.adj option), the software requires the rms package, distributed by Frank Harrell [5]. If the user does not have these packages installed, then it will not be possible to use the x.adj option. However, the unadjusted scores for an ordinal trait (using the default option x.adj=NA) do not require these packageses. Check the list of your local packages in the list shown from entering library() in your prompt.

5.4 Haplotype Scores, Adjusted for Covariates

To adjust for covariates in haplo.score, first set up a matrix of covariates from the example data. For example, use a column for male (1 if male; 0 if female), and a second column for age. Then pass the matrix to haplo.score using parameter x.adj. The results change slightly in this example.

```
> # score w/gaussian, adjusted by covariates
> x.ma <- cbind(male, age)</pre>
```

```
> score.gaus.adj <- haplo.score(resp, geno, trait.type="gaussian",
                  x.adj = x.ma, min.count=5,
+
                  locus.label=label, simulate=FALSE)
+
> print(score.gaus.adj, nlines=10)
                 _____
                Haplotype Effect Model: additive
                             Global Score Statistics
_____
global-stat = 31.02908, df = 18, p-val = 0.02857
_____
                   Haplotype-specific Scores
_____
    DQB DRB B Hap-Freq Hap-Score p-val
         8 0.10408 -2.4097
[1,] 21
                         0.01597
       3
[2,] 31
       4
         44 0.02849 -2.25293 0.02426
[3,] 51
       1
         44 0.01731 -0.98763 0.32333
[4,] 63
      13 44 0.01606 -0.83952 0.40118
[5,] 63
      2
         7 0.01333 -0.48483 0.6278
[6,] 32
      4
         60 0.0306
                  -0.46476 0.64211
      7
[7,] 21
         44 0.02332 -0.41249 0.67998
[8,] 62
      2
        44 0.01367
                  -0.26443 0.79145
[9,] 62
       2
         18 0.01545
                  -0.20425
                         0.83816
[10,] 51
      1
         27 0.01505 0.02243
                         0.9821
```

5.5 Plots and Haplotype Labels

A convenient way to view results from haplo.score is a plot of the haplotype frequencies (Hap-Freq) versus the haplotype score statistics (Hap-Score). This plot, and the syntax for creating it, are shown in Figure 1.

Some points on the plot may be of interest. To identify individual points on the plot, use *locator.haplo(score.gaus)*, which is similar to *locator()*. Use the mouse to select points on the plot. After points are chosen, click on the middle mouse button, and the points are labeled with their haplotype labels. Note, in constructing Figure 1, we had to define which points to label, and then assign labels in the same way as done within the *locator.haplo* function.



Figure 1: Haplotype Statistics: Score vs. Frequency, Quantitative Response

5.6 Skipping Rare Haplotypes

For the haplo.score, the skip.haplo and min.count parameters control which rare haplotypes are pooled into a common group. The min.count parameter is a recent addition to haplo.score, yet it does the same task as skip.haplo and is the same idea as haplo.min.count used in haplo.glm.control for haplo.glm. As a guideline, you may wish to set min.count to calculate scores for haplotypes with expected haplotype counts of 5 or greater in the sample. We concentrate on this expected count because it adjusts to the size of the input data. If N is the number of subjects and f the haplotype frequency, then the expected haplotype count is $count = 2 \times N \times f$. Alternatively, you can choose $skip.haplo = \frac{count}{2 \times N}$.

In the following example we try a different cut-off than before, min.count=10, which corresponds to skip.haplo of $10 \div (2 \times 220) = .045$. In the output, see that the global statistic, degrees of freedom, and p-value change because of the fewer haplotypes, while the haplotype-specific scores do not change.

```
> # increase skip.haplo, expected hap counts = 10
> score.gaus.min10 <- haplo.score(resp, geno, trait.type="gaussian",</pre>
+
                     x.adj = NA, min.count=10,
                     locus.label=label, simulate=FALSE)
> print(score.gaus.min10)
                  Haplotype Effect Model: additive
                _____
                     Global Score Statistics
 _____
global-stat = 20.66451, df = 7, p-val = 0.0043
                      -----
                    Haplotype-specific Scores
_____
                     _____
   DQB DRB B Hap-Freq Hap-Score p-val
[1,] 21
      З
         8 0.10408 -2.39631 0.01656
[2,] 31
         44 0.02849 -2.24273 0.02491
      4
[3,] 32
      4
         60 0.0306 -0.46606 0.64118
[4,] 21
      7
         44 0.02332 -0.41942 0.67491
[5,] 51
         35 0.03018 0.69696
      1
                          0.48583
[6,] 32 4
         62 0.02349 2.37619
                          0.01749
```

5.7 Score Statistic Dependencies: the eps.svd parameter

7 0.05098 2.39795

[7,] 62 2

The global score test is calculated using the vector of scores and the generalized inverse of their variance-covariance matrix, performed by the *Ginv* function. This function determines the rank

0.01649

of the variance matrix by its singular value decomposition, and an epsilon value is used as the cut-off for small singular values. If all of the haplotypes in the sample are scored, then there is dependence between them and the variance matrix is not of full rank. However, it is more often the case that one or more rare haplotypes are not scored because of low frequency. It is not clear how strong the dependencies are between the remaining score statistics, and likewise, there is disparity in calculating the rank of the variance matrix. For these instances we give the user control over the epsilon parameter for haplo.score with eps.svd.

We have seen instances where the global score test had a very significant p-value, but none of the haplotype-specific scores showed strong association. In such instances, we found the default epsilon value in *Ginv* was incorrectly considering the variance matrix as having full rank, and the misleading global score test was corrected when we increased epsilon for *Ginv*. We now set the default for *eps.svd* at 1e - 5, which seems to perform better in the troublesome circumstances than the previous default of 1e - 6.

5.8 Haplotype Model Effect

A recent addition to haplo.score is the ability to select non-additive effects to score haplotypes. The possible effects for haplotypes are additive, dominant, and recessive. Under recessive effects, fewer haplotypes may be scored, because subjects are required to be homozygous for haplotypes. Furthermore, there would have to be *min.count* such persons in the sample to have the recessive effect scored. Therefore, a recessive model should only be used on samples with common haplotypes. In the example below with the gaussian response, set the haplotype effect to dominant using parameter haplo.effect = "dominant". Notice the results change slightly compared to the score.gaus.add results above.

```
> # score w/gaussian, dominant effect
>
> score.gaus.dom <- haplo.score(resp, geno, trait.type="gaussian",
+
                    x.adj=NA, min.count=5,
                    haplo.effect="dominant", locus.label=label,
+
                    simulate=FALSE)
> print(score.gaus.dom, nlines=10)
                             _____
                  Haplotype Effect Model: dominant
                      Global Score Statistics
    _____
global-stat = 29.56133, df = 18, p-val = 0.04194
                                     _____
                     Haplotype-specific Scores
```

	DQB	DRB	В	Hap-Freq	Hap-Score	p-val
[1,]	21	3	8	0.10408	-2.23872	0.02517
[2,]	31	4	44	0.02849	-2.13233	0.03298
[3,]	51	1	44	0.01731	-0.99357	0.32043
[4,]	63	13	44	0.01606	-0.84453	0.39837
[5,]	63	2	7	0.01333	-0.50736	0.6119
[6,]	32	4	60	0.0306	-0.46606	0.64118
[7,]	21	7	44	0.02332	-0.41942	0.67491
[8,]	62	2	44	0.01367	-0.26221	0.79316
[9,]	62	2	18	0.01545	-0.21493	0.82982
[10,]	51	1	27	0.01505	0.01539	0.98772

5.9 Simulation p-values

When simulate=TRUE, haplo.score gives simulated p-values. Simulated haplotype score statistics are the re-calculated score statistics from a permuted re-ordering of the trait and covariates and the original ordering of the genotype matrix. The simulated p-value for the global score statistic (Global sim. p-val) is the number of times the simulated global score statistic exceeds the observed, divided by the total number of simulations. Likewise, simulated p-value for the maximum score statistic (Max-stat sim. p-val) is the number of times the simulated maximum haplotype score statistic exceeds the observed maximum score statistic, divided by the total number of simulations. The maximum score statistic is the maximum of the square of the haplotype-specific score statistics, which has an unknown distribution, so its significance can only be given by the simulated p-value. Intuitively, if only one or two haplotypes are associated with the trait, the maximum score statistic should have greater power to detect association than the global statistic.

The score.sim.control function manages control parameters for simulations. The haplo.score function employs the simulation p-value precision criteria of Besag and Clifford[6]. These criteria ensure that the simulated p-values for both the global and the maximum score statistics are precise for small p-values. The algorithm performs a user-defined minimum number of permutations (min.sim) to guarantee sufficient precision for the simulated p-values for score statistics of individual haplotypes. Permutations beyond this minimum are then conducted until the sample standard errors for simulated p-values for both the global-stat and max-stat score statistics are less than a threshold (p.threshold * p-value). The default value for p.threshold= $\frac{1}{4}$ provides a two-sided 95% confidence interval for the p-value with a width that is approximately as wide as the p-value itself. Effectively, simulations are more precise for smaller p-values. The following example illustrates computation of simulation p-values with min.sim=1000.

```
> # simulations when binary response
> score.bin.sim <- haplo.score(y.bin, geno, trait.type="binomial",
+ x.adj = NA, locus.label=label, min.count=5,
+ simulate=TRUE, sim.control = score.sim.control() )
> print(score.bin.sim)
```

Haplotype Effect Model: additive

-------Global Score Statistics _____ global-stat = 33.70125, df = 18, p-val = 0.01371 _____ Global Simulation p-value Results _____ _____ Global sim. p-val = 0.0095 Max-Stat sim. p-val = 0.00563 Number of Simulations, Global: 2842 , Max-Stat: 2842 _____ Haplotype-specific Scores _____ _____ DQB DRB B Hap-Freq Hap-Score p-val sim p-val [1,] 62 2 7 0.05098 -2.19387 0.02824 0.02991 [2,] 51 1 35 0.03018 -1.58421 0.11315 0.13863 [3,] 63 13 7 0.01655 -1.56008 0.11874 0.19177 [4,] 21 7 7 0.01246 -1.47495 0.14023 0.15588 7 0.01678 -1.00091 0.31687 0.25123 [5,] 32 4 [6,] 32 4 62 0.02349 -0.6799 0.49657 0.47467 [7,] 51 27 0.01505 -0.66509 0.50599 0.63089 1 [8,] 31 11 35 0.01754 -0.5838 0.55936 0.6506 [9,] 31 11 51 0.01137 -0.43721 0.66196 0.91872 [10,] 51 44 0.01731 0.00826 0.99341 1 1 [11,] 32 60 0.0306 0.03181 0.97462 0.95074 4 [12,] 62 2 44 0.01367 0.16582 0.8683 0.91872 [13,] 63 13 44 0.01606 0.22059 0.82541 0.7266 [14,] 63 2 7 0.01333 0.2982 0.76555 0.89163 [15,] 62 2 18 0.01545 0.78854 0.43038 0.6608 [16,] 21 7 44 0.02332 0.84562 0.39776 0.39796 [17,] 31 4 44 0.02849 2.50767 0.01215 0.01161 [18,] 21 8 0.10408 3.77763 3 0.00016 0.00035

6 Regression Models: haplo.glm

The haplo.glm function computes the regression of a trait on haplotypes, and possibly other covariates and their interactions with haplotypes. We currently support the gaussian, binomial, and Poisson families of traits with their canonical link functions. The effects of haplotypes on the link function can be modeled as either additive, dominant (heterozygotes and homozygotes for a particular haplotype assumed to have equivalent effects), or recessive (homozygotes of a particular haplotype considered to have an alternative effect on the trait). The basis of the algorithm is a two-step iteration process; the posterior probabilities of pairs of haplotypes per subject are used as weights to update the regression coefficients, and the regression coefficients are used to update the haplotype posterior probabilities. See Lake et al.[7] for details.

6.1 New and Updated Methods for haplo.glm

We initially wrote haplo.glm without much effort of making it work with R's glm class methods. We have now refined the haplo.glm class to look and act as much like a glm object as possible with methods defined specifically for the haplo.glm class. We provide print and summary methods that make use of the corresponding methods for glm and then add extra information for the haplotypes and their frequencies. Furthermore, we have defined for the haplo.glm class some of the standard methods for regression fits, including residuals, fitted.values, vcov, and anova. We describe the challenges that haplotype regression presents with these methods in section 7.

6.2 Preparing the data.frame for haplo.glm

A critical distinction between haplo.glm and all other functions in Haplo Stats is that the definition of the regression model follows the S/R formula standard (see lm or glm). So, a data.frame must be defined, and this object must contain the trait and other optional covariates, plus a special kind of genotype matrix (geno.glm for this example) that contains the genotypes of the marker loci. We require the genotype matrix to be prepared using setupGeno(), which handles character, numeric, or factor alleles, and keeps the columns of the genotype matrix as a single unit when inserting into (and extracting from) a data.frame. The setupGeno function recodes all missing genotype value codes given by miss.val to NA, and also recodes alleles to integer values. The original allele codes are preserved within an attribute of geno.glm, and are utilized within haplo.glm. The returned object has class model.matrix, and it can be included in a data.frame to be used in haplo.glm. In the example below we prepare a genotype matrix, geno.glm, and create a data.frame object, glm.data, for use in haplo.glm.

```
> # set up data for haplo.glm, include geno.glm,
> # covariates age and male, and responses resp and y.bin
> geno <- hla.demo[,c(17,18,21:24)]
> geno.glm <- setupGeno(geno, miss.val=c(0,NA), locus.label=label)
> attributes(geno.glm)
$dim
[1] 220 6
$dimnames
$dimnames[[1]]
NULL
$dimnames[[2]]
[1] "DQB.a1" "DQB.a2" "DRB.a1" "DRB.a2" "B.a1" "B.a2"
```

```
$class
[1] "model.matrix"
$unique.alleles
$unique.alleles[[1]]
[1] "21" "31" "32" "33" "42" "51" "52" "53" "61" "62" "63" "64"
$unique.alleles[[2]]
[1] "1" "2" "3" "4" "7" "8" "9" "10" "11" "13" "14"
$unique.alleles[[3]]
[1] "7" "8" "13" "14" "18" "27" "35" "37" "38" "39" "41" "42" "44" "45" "46"
[16] "47" "48" "49" "50" "51" "52" "55" "56" "57" "58" "60" "61" "62" "63" "70"
> y.bin <- 1*(resp.cat=="low")
> glm.data <- data.frame(geno.glm, age=age, male=male, y=resp, y.bin=y.bin)</pre>
```

6.3 Rare Haplotypes

The issue of deciding which haplotypes to use for association is critical in *haplo.glm*. By default it will model a rare haplotype effect so that the effects of other haplotypes are in reference to the baseline effect of the one common happlotype. The rules for choosing haplotypes to be modeled in *haplo.glm* are similar to the rules in *haplo.score*: by a minimum frequency or a minimum expected count in the sample.

Two control parameters in haplo.glm.control may be used to control this setting: haplo.freq.min may be set to a selected minimum haplotype frequency, and haplo.min.count may be set to select the cut-off for minimum expected haplotype count in the sample. The default minimum frequency cut-off in haplo.glm is set to 0.01. More discussion on rare haplotypes takes place in section 6.7.4.

6.4 Regression for a Quantitative Trait

The following illustrates how to fit a regression of a quantitative trait y on the haplotypes estimated from the geno.glm matrix, and the covariate male. For na.action, we use na.geno.keep, which keeps a subject with missing values in the genotype matrix if they are not missing all alleles, but removes subjects with missing values (NA) in either the response or covariate.

```
> # glm fit with haplotypes, additive gender covariate on gaussian response
> fit.gaus <- haplo.glm(y ~ male + geno.glm, family=gaussian, data=glm.data,
+ na.action="na.geno.keep", locus.label = label, x=TRUE,
+ control=haplo.glm.control(haplo.freq.min=.02))
> summary(fit.gaus)
Call:
haplo.glm(formula = y ~ male + geno.glm, family = gaussian, data = glm.data,
na.action = "na.geno.keep", locus.label = label, control = haplo.glm.control(haplo.freq.
x = TRUE)
```

Deviance Residuals: Min 1Q Median ЗQ Max -2.46945 -0.92052 -0.06533 0.94874 2.37199 Coefficients: t.stat pval coef se (Intercept) 1.06436 0.34283 3.10464 0.002 male 0.09735 0.15521 0.62723 0.531 geno.glm.17 0.28022 0.43549 0.64346 0.521 geno.glm.34 -0.31713 0.34342 -0.92342 0.357 geno.glm.77 0.22167 0.36126 0.61360 0.540 geno.glm.78 0.38382 2.97390 0.003 1.14144 geno.glm.100 0.55557 0.36427 1.52517 0.129 geno.glm.138 0.98229 0.30329 3.23875 0.001 geno.glm.rare 0.39765 0.18191 2.18591 0.030

(Dispersion parameter for gaussian family taken to be 1.269581)

Null deviance: 297.01 on 219 degrees of freedom Residual deviance: 267.88 on 211 degrees of freedom AIC: 687.65

Number of Fisher Scoring iterations: 268

Haplotypes:

	DQB	DRB	В	hap.freq
geno.glm.17	21	7	44	0.02291
geno.glm.34	31	4	44	0.02858
geno.glm.77	32	4	60	0.03022
geno.glm.78	32	4	62	0.02390
geno.glm.100	51	1	35	0.03008
geno.glm.138	62	2	7	0.05023
<pre>geno.glm.rare</pre>	*	*	*	0.71000
haplo.base	21	3	8	0.10409

Explanation of Results

The new summary function for haplo.glm shows much the same information as summary for glm objects with the extra table for the haplotype frequencies. The above table for *Coefficients* lists the estimated regression coefficients (*coef*), standard errors (*se*), the corresponding t-statistics (*t.stat*), and p-values (*pval*). The labels for haplotype coefficients are a concatenation of the name of the genotype matrix (*geno.glm*) and unique haplotype codes assigned within haplo.glm. The haplotypes corresponding to these haplotype codes are listed in the Haplotypes table, along with the estimates of the haplotype frequencies (*hap.freq*). The rare haplotypes, those with expected counts less than haplo.min.count=5 (equivalent to having frequencies less than

haplo.freq.min = 0.011363636363636364) in the above example), are pooled into a single category labeled geno.glm.rare. The haplotype chosen as the baseline category for the design matrix (most frequent haplotype is the default) is labeled as haplo.base; more information on the baseline may be found in section 6.7.2.

6.5 Fitting Haplotype x Covariate Interactions

Interactions are fit by the standard S-language model syntax, using a '*' in the model formula to indicate main effects and interactions. Some other formula constructs are not supported, so use the formula parameter with caution. Below is an example of modeling the interaction of *male* and the haplotypes. Because more terms will be estimated in this case, we limit how many haplotypes will be included by increasing *haplo.min.count* to 10.

```
> # glm fit haplotypes with covariate interaction
> fit.inter <- haplo.glm(formula = y ~ male * geno.glm,</pre>
+
                     family = gaussian, data=glm.data,
                     na.action="na.geno.keep",
+
+
                     locus.label = label,
                     control = haplo.glm.control(haplo.min.count = 10))
+
> summary(fit.inter)
Call:
haplo.glm(formula = y ~ male * geno.glm, family = gaussian, data = glm.data,
   na.action = "na.geno.keep", locus.label = label, control = haplo.glm.control(haplo.min.c
Deviance Residuals:
    Min
                1Q
                      Median
                                    ЗQ
                                              Max
-2.23387
         -0.90661 -0.05953
                               0.96140
                                         2.48859
Coefficients:
                       coef
                                  se
                                       t.stat pval
                    0.97536
                             0.52268
                                     1.86607 0.063
(Intercept)
male
                    0.25806
                             0.67351
                                      0.38315 0.702
geno.glm.17
                             0.54544 0.26479 0.791
                    0.14443
                             0.66773 -0.25700 0.797
geno.glm.34
                   -0.17161
geno.glm.77
                    0.80523
                             0.64951 1.23975 0.216
geno.glm.78
                    0.49557
                             0.56574 0.87596 0.382
geno.glm.100
                    0.52310
                             0.48067
                                      1.08828 0.278
geno.glm.138
                    1.15704
                             0.42325 2.73371 0.007
geno.glm.rare
                    0.45547
                             0.28721
                                      1.58587 0.114
male:geno.glm.17
                    0.50872
                             0.87531 0.58119 0.562
male:geno.glm.34
                   -0.28137
                             0.78570 -0.35812 0.721
male:geno.glm.77
                   -0.90084
                             0.79114 -1.13865 0.256
male:geno.glm.78
                    1.26376
                             0.77131
                                      1.63846 0.103
male:geno.glm.100
                             0.77470
                                      0.06549 0.948
                    0.05074
male:geno.glm.138
                   -0.44587
                             0.61903 -0.72027 0.472
```

male:geno.glm.rare -0.09787 0.37197 -0.26312 0.793 (Dispersion parameter for gaussian family taken to be 1.27362) Null deviance: 297.01 on 219 degrees of freedom Residual deviance: 259.82 on 204 degrees of freedom AIC: 694.93 Number of Fisher Scoring iterations: 120 Haplotypes: DQB DRB B hap.freq geno.glm.17 21 7 44 0.02346 geno.glm.34 31 4 44 0.02845 geno.glm.77 32 4 60 0.03060 geno.glm.78 32 4 62 0.02413 0.03013 geno.glm.100 51 1 35 geno.glm.138 62 2 7 0.05049 geno.glm.rare * * * 0.70863 haplo.base 21 3 8 0.10410

Explanation of Results

The listed results are as explained under section 6.4. The main difference is that the interaction coefficients are labeled as a concatenation of the covariate (*male* in this example) and the name of the haplotype, as described above. In addition, estimates may differ because the model has changed.

6.6 Regression for a Binomial Trait

Next we illustrate the fitting of a binomial trait with the same genotype matrix and covariate.

```
> # gender and haplotypes fit on binary response,
> # return model matrix
> fit.bin <- haplo.glm(y.bin ~ male + geno.glm, family = binomial,
+ data=glm.data, na.action = "na.geno.keep",
+ locus.label=label,
+ control = haplo.glm.control(haplo.min.count=10))
> summary(fit.bin)
Call:
haplo.glm(formula = y.bin ~ male + geno.glm, family = binomial,
data = glm.data, na.action = "na.geno.keep", locus.label = label,
control = haplo.glm.control(haplo.min.count = 10))
```

Deviance Residuals:

Min 1Q Median 3Q Max -1.5559 -0.7996 -0.6473 1.0591 2.4348

Coefficients:

	coef	se	t.stat	pval
(Intercept)	1.5457	0.6547	2.3610	0.019
male	-0.4802	0.3308	-1.4518	0.148
geno.glm.17	-0.7227	0.8011	-0.9022	0.368
geno.glm.34	0.3641	0.6798	0.5356	0.593
geno.glm.77	-0.9884	0.7328	-1.3489	0.179
geno.glm.78	-1.4093	0.8543	-1.6496	0.101
geno.glm.100	-2.5907	1.1278	-2.2971	0.023
geno.glm.138	-2.7156	0.8524	-3.1860	0.002
geno.glm.rare	-1.2610	0.3537	-3.5647	0.000

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 263.50 on 219 degrees of freedom Residual deviance: 233.46 on 211 degrees of freedom AIC: 251.11

Number of Fisher Scoring iterations: 61

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	DQB	DRB	В	hap.freq
geno.glm.17	21	7	44	0.02303
geno.glm.34	31	4	44	0.02843
geno.glm.77	32	4	60	0.03057
geno.glm.78	32	4	62	0.02354
geno.glm.100	51	1	35	0.02977
geno.glm.138	62	2	7	0.05181
geno.glm.rare	*	*	*	0.70880
haplo.base	21	3	8	0.10405

Explanation of Results

The underlying methods for *haplo.glm* are based on a prospective likelihood. Normally, this type of likelihood works well for case-control studies with standard covariates. For ambiguous haplotypes, however, one needs to be careful when interpreting the results from fitting *haplo.glm* to case-control data. Because cases are over-sampled, relative to the population prevalence (or incidence, for incident cases), haplotypes associated with disease will be over-represented in the case sample, and so estimates of haplotype frequencies will be biased. Positively associated haplotypes will have haplotype frequency estimates that are higher than the population haplotype frequency. To avoid this problem, one can weight each subject. The weights for the cases should be the population prevalence, and the weights for controls should be 1 (assuming the disease is rare in

the population, and controls are representative of the general population). See Stram et al.[8] for background on using weights, and see the help file for *haplo.glm* for how to implement weights.

The estimated regression coefficients for case-control studies can be biased by either a large amount of haplotype ambiguity and mis-specified weights, or by departures from Hardy-Weinberg Equilibrium of the haplotypes in the pool of cases and controls. Generally, the bias is small, but tends to be towards the null of no association. See Stram et al. [8] and Epstein and Satten [9] for further details.

6.6.1 Caution on Rare Haplotypes with Binomial Response

If a rare haplotype occurs only in cases or only in controls, the fitted values would go to 0 or 1, where R would issue a warning. Also, the coefficient estimate for that haplotype would go to positive or negative infinity, If the default *haplo.min.count=5* were used above, this warning would appear. To keep this from occuring in other model fits, increase the minimum count or minimum frequency.

6.7 Control Parameters

Additional parameters are handled using *control*, which is a list of parameters providing additional functionality in *haplo.glm*. This list is set up by the function *haplo.glm.control*. See the help file (*help(haplo.glm.control*)) for a full list of control parameters, with details of their usage. Some of the options are described here.

6.7.1 Controlling Genetic Models: haplo.effect

The haplo effect control parameter for haplo glm instructs whether the haplotype effects are fit as additive, dominant, or recessive. That is, haplo effect determines whether the covariate (x) coding of haplotypes follows the values in Table 1 for each effect type. Heterozygous means a subject has one copy of a particular haplotype, and homozygous means a subject has two copies of a particular haplotype.

 Table 1: Coding haplotype covariates in a model matrix

Hap - Pair	additive	dominant	recessive
Heterozygous	1	1	0
Homozygous	2	1	1

Note that in a recessive model, the haplotype effects are estimated only from subjects who are homozygous for a haplotype. Some of the haplotypes which meet the *haplo.freq.min* and *haplo.count.min* cut-offs may occur as homozygous in only a few of the subjects. As stated in 5.8, recessive models should be used when the region has multiple common haplotypes.

The default haplo effect is additive, whereas the example below illustrates the fit of a dominant effect of haplotypes for the gaussian trait with the gender covariate.

```
> # control dominant effect of haplotypes (haplo.effect)
> # by using haplo.glm.control
> fit.dom <- haplo.glm(y ~ male + geno.glm, family = gaussian,</pre>
        data = glm.data, na.action = "na.geno.keep",
+
+
        locus.label = label,
        control = haplo.glm.control(haplo.effect='dominant',
+
+
        haplo.min.count=8))
> summary(fit.dom)
Call:
haplo.glm(formula = y ~ male + geno.glm, family = gaussian, data = glm.data,
   na.action = "na.geno.keep", locus.label = label, control = haplo.glm.control(haplo.effect
       haplo.min.count = 8))
Deviance Residuals:
    Min
               1Q
                     Median
                                   ЗQ
                                            Max
-2.48099 -1.01196
                    0.01035 1.00557
                                        2.48801
Coefficients:
                               t.stat pval
                 coef
                            se
(Intercept)
              1.64935 0.37350 4.41593 0.000
              0.07969 0.15726 0.50673 0.613
male
geno.glm.17
             -0.06035 0.42317 -0.14262 0.887
geno.glm.34 -0.66499 0.36392 -1.82731 0.069
geno.glm.77
             -0.07339 0.34665 -0.21171 0.833
geno.glm.78
             0.85369 0.36421 2.34394 0.020
geno.glm.100
              0.24697 0.34561 0.71458 0.476
              0.67295 0.28163 2.38944 0.018
geno.glm.138
geno.glm.rare 0.11195 0.34006 0.32922 0.742
(Dispersion parameter for gaussian family taken to be 1.300586)
    Null deviance: 297.01 on 219 degrees of freedom
Residual deviance: 274.42 on 211 degrees of freedom
AIC: 692.96
Number of Fisher Scoring iterations: 91
Haplotypes:
             DQB DRB B hap.freq
              21
                   7 44 0.02297
geno.glm.17
geno.glm.34
              31
                   4 44 0.02855
              32 4 60 0.03019
geno.glm.77
geno.glm.78
              32 4 62 0.02391
geno.glm.100
              51 1 35 0.03003
                                       34
```

geno.glm.138	62	2	7	0.05023
geno.glm.rare	*	*	*	0.71003
haplo.base	21	3	8	0.10408

6.7.2 Selecting the Baseline Haplotype

The haplotype chosen for the baseline in the model is the one with the highest frequency. Sometimes the most frequent haplotype may be an at-risk haplotype, and so the measure of its effect is desired. To specify a more appropriate haplotype as the baseline in the binomial example, choose from the list of other common haplotypes, *fit.bin\$haplo.common*. To specify an alternative baseline, such as haplotype 77, use the control parameter *haplo.base* and haplotype code, as in the example below.

```
> # control baseline selection, perform the same exact run as fit.bin,
> # but different baseline by using haplo.base chosen from haplo.common
> fit.bin$haplo.common
[1] 17 34 77 78 100 138
> fit.bin$haplo.freq.init[fit.bin$haplo.common]
[1] 0.02332031 0.02848720 0.03060053 0.02349463 0.03018431 0.05097906
> fit.bin.base77 <- haplo.glm(y.bin ~ male + geno.glm, family = binomial,</pre>
         data = glm.data, na.action = "na.geno.keep",
+
+
         locus.label = label,
+
         control = haplo.glm.control(haplo.base=77,
         haplo.min.count=8))
+
> summary(fit.bin.base77)
Call:
haplo.glm(formula = y.bin ~ male + geno.glm, family = binomial,
    data = glm.data, na.action = "na.geno.keep", locus.label = label,
    control = haplo.glm.control(haplo.base = 77, haplo.min.count = 8))
Deviance Residuals:
    Min
              1Q
                   Median
                                ЗQ
                                        Max
-1.5559 -0.7996 -0.6473
                            1.0591
                                     2.4348
Coefficients:
                           se t.stat pval
                 coef
(Intercept)
              -0.4311 1.3586 -0.3173 0.751
              -0.4802 0.3308 -1.4518 0.148
male
               0.9884 0.7328 1.3489 0.179
geno.glm.4
geno.glm.17
               0.2657
                       1.0254 0.2591 0.796
                       0.9223 1.4665 0.144
geno.glm.34
               1.3525
geno.glm.78
              -0.4209
                       1.0430 -0.4035 0.687
geno.glm.100
             -1.6023 1.3007 -1.2319 0.219
```

geno.glm.138 -1.7273 1.0321 -1.6736 0.096 geno.glm.rare -0.2726 0.6834 -0.3989 0.690 (Dispersion parameter for binomial family taken to be 1) Null deviance: 263.50 on 219 degrees of freedom Residual deviance: 233.46 on 211 degrees of freedom AIC: 251.11 Number of Fisher Scoring iterations: 61 Haplotypes: DQB DRB B hap.freq geno.glm.4 21 3 8 0.10405 geno.glm.17 21 7 44 0.02303 geno.glm.34 31 4 44 0.02843 geno.glm.78 32 4 62 0.02354 geno.glm.100 51 1 35 0.02977 geno.glm.138 62 2 7 0.05181 geno.glm.rare * * * 0.70880 32 haplo.base 4 60 0.03057

Explanation of Results

The above model has the same haplotypes as *fit.bin*, except haplotype 4, the old baseline, now has an effect estimate while haplotype 77 is the new baseline. Due to randomness in the starting values of the haplotype frequency estimation, different runs of *haplo.glm* may result in a different set of haplotypes meeting the minimum counts requirement for being modeled. Therefore, once you have arrived at a suitable model, and you wish to modify it by changing baseline and/or effects, you can make results consistent by controlling the randomness using *set.seed*, as described in section 2.4. In this document, we use the same seed before making *fit.bin* and *fit.bin.base77*.

6.7.3 Rank of Information Matrix and eps.svd (NEW)

Similar to recent additions to haplo.score in section 5.7, we give the user control over the epsilon parameter determining the number of singular values when determining the rank of the information matrix in haplo.glm. Finding the generalized inverse of this matrix can be problematic when either the response variable or a covariate has a large variance and is not scaled before passed to haplo.glm. The rank of the information matrix is determined by the number of non-zero singular values a small cutoff, epsilon. When the singular values for the coefficients are on a larger numeric scale than those for the haplotype frequencies, the generalized inverse may incorrectly determine the information matrix is not of full rank. Therefore, we allow the user to specify the epsilon as eps.svd in the control parameters for haplo.glm. A simpler fix, which we strongly suggest, is for the user to pre-scale any continuous responses or covariates with a large variance.

Here we demonstrate what happens when we increase the variance of a gaussian response by
2500. We see that the coefficients are all highly significant and the rank of the information matrix is much smaller than the scaled gaussian fit.

```
> glm.data$ybig <- glm.data$y*50</pre>
> fit.gausbig <- haplo.glm(formula = ybig ~ male + geno.glm, family = gaussian,</pre>
           data = glm.data, na.action = "na.geno.keep", locus.label = label,
+
           control = haplo.glm.control(haplo.freq.min = 0.02), x = TRUE)
+
> summary(fit.gausbig)
Call:
haplo.glm(formula = ybig ~ male + geno.glm, family = gaussian,
    data = glm.data, na.action = "na.geno.keep", locus.label = label,
    control = haplo.glm.control(haplo.freq.min = 0.02), x = TRUE)
Deviance Residuals:
                     Median
                                   ЗQ
    Min
               1Q
                                            Max
-123.472
         -46.026
                     -3.267
                               47.450
                                        118.550
Coefficients:
                 coef
                            se
                               t.stat pval
                        1.7343 30.6849 0.000
(Intercept)
              53.2180
               4.8675 6.0042 0.8107 0.418
male
geno.glm.17
             14.0111 0.2579 54.3195 0.000
geno.glm.34
             -15.8563 1.0033 -15.8044 0.000
geno.glm.77
              11.0835 0.9978 11.1078 0.000
geno.glm.78
              57.0720
                        0.3855 148.0436 0.000
geno.glm.100
              27.7784 0.3228 86.0564 0.000
geno.glm.138
              49.1143
                        1.0334 47.5256 0.000
                       3.4583 5.7493 0.000
geno.glm.rare
              19.8824
(Dispersion parameter for gaussian family taken to be 3173.952)
    Null deviance: 742530 on 219 degrees of freedom
Residual deviance: 669704 on 211 degrees of freedom
AIC: 2408.9
Number of Fisher Scoring iterations: 268
Haplotypes:
             DQB DRB B hap.freq
geno.glm.17
              21
                   7 44 0.02291
geno.glm.34
              31
                   4 44 0.02858
geno.glm.77
              32
                   4 60 0.03022
geno.glm.78
              32
                   4 62 0.02390
geno.glm.100
              51 1 35 0.03008
                                       37
```

geno.glm.138 62 2 7 0.05023 geno.glm.rare * * * 0.71000 haplo.base 21 3 8 0.10409

> fit.gausbig\$rank.info

[1] 175

> fit.gaus\$rank.info

[1] 182

Now we set a smaller value for the *eps.svd* control parameter and find the fit matches the original Gaussian fit.

```
> fit.gausbig.eps <- haplo.glm(formula = ybig ~ male + geno.glm, family = gaussian,</pre>
            data = glm.data, na.action = "na.geno.keep", locus.label = label,
+
            control = haplo.glm.control(eps.svd=1e-10, haplo.freq.min = 0.02), x = TRUE)
+
> summary(fit.gausbig.eps)
Call:
haplo.glm(formula = ybig ~ male + geno.glm, family = gaussian,
    data = glm.data, na.action = "na.geno.keep", locus.label = label,
    control = haplo.glm.control(eps.svd = 1e-10, haplo.freq.min = 0.02),
    x = TRUE)
Deviance Residuals:
    Min
                     Median
                                   3Q
                                             Max
               1Q
-123.472
          -46.026
                     -3.267
                                47.450
                                         118.550
Coefficients:
                                 t.stat pval
                 coef
                            se
                                 3.1046 0.002
(Intercept)
              53.2180 17.1414
male
               4.8675
                       7.7603
                                 0.6272 0.531
geno.glm.17
              14.0111
                       21.7745
                                 0.6435 0.521
             -15.8563 17.1712 -0.9234 0.357
geno.glm.34
geno.glm.77
              11.0835 18.0631 0.6136 0.540
geno.glm.78
              57.0720 19.1910
                                 2.9739 0.003
              27.7784 18.2133
                                 1.5252 0.129
geno.glm.100
geno.glm.138
               49.1143 15.1646
                                 3.2387 0.001
geno.glm.rare
              19.8824
                       9.0957
                                 2.1859 0.030
(Dispersion parameter for gaussian family taken to be 3173.952)
```

```
Null deviance: 742530 on 219 degrees of freedom
Residual deviance: 669704 on 211 degrees of freedom
AIC: 2408.9
```

Number of Fisher Scoring iterations: 268

Haplotypes:				
	DQB	DRB	В	hap.freq
geno.glm.17	21	7	44	0.02291
geno.glm.34	31	4	44	0.02858
geno.glm.77	32	4	60	0.03022
geno.glm.78	32	4	62	0.02390
geno.glm.100	51	1	35	0.03008
geno.glm.138	62	2	7	0.05023
<pre>geno.glm.rare</pre>	*	*	*	0.71000
haplo.base	21	3	8	0.10409

> fit.gausbig.eps\$rank.info

[1] 182

6.7.4 Rare Haplotypes and haplo.min.info

Another notable control parameter is the minimum frequency for a rare haplotype to be included in the calculations for standard error (se) of the coefficients, or *haplo.min.info*. The default value is 0.001, which means that haplotypes with frequency less than that will be part of the rare haplotype coefficient estimate, but it will not be used in the standard error calculation.

The following example demonstrates a possible result when dealing with the rare haplotype effect. We show with the hla genotype data one consequence for when this occurs. However, we make it happen by setting *haplo.freq.min* equal to *haplo.min.info*, which we advise strongly against in your analyses.

Min	1Q	Median	ЗQ	Max
-1.4558	-0.7238	-0.7238	1.0382	2.3083

Coefficients: coef t.stat pval se (Intercept) 1.2409 1.3238 0.9374 0.350 geno.glm.17 -0.6068 1.5630 -0.3882 0.698 geno.glm.34 0.2516 0.802 0.3189 1.2678 geno.glm.77 -1.07191.3666 -0.7844 0.434 4.3659 -0.3113 0.756 geno.glm.78 -1.3593geno.glm.100 -2.39842.0878 -1.1487 0.252 geno.glm.138 -2.60961.5043 -1.7347 0.084 geno.glm.rare -1.2233 NA NA NA (Dispersion parameter for binomial family taken to be 1) Null deviance: 263.50 on 219 degrees of freedom Residual deviance: 235.59 on 212 degrees of freedom AIC: 251.24 Number of Fisher Scoring iterations: 71 Haplotypes: DQB DRB B hap.freq 0.02302 geno.glm.17 21 7 44 geno.glm.34 31 4 44 0.02841 4 60 geno.glm.77 32 0.03058 geno.glm.78 32 4 62 0.02353 geno.glm.100 51 1 35 0.02980 geno.glm.138 62 2 7 0.05190 geno.glm.rare * * 0.70872 * haplo.base 21 3 8 0.10403 **Explanation of Results**

The above results show the standard error for the rare haplotype coefficient is "NaN", or "Not a Number" in R, which is a consequence of having most, or all, of the rare haplotypes discarded for the standard error estimate. In other datasets there may be only a few haplotypes between haplo.min.info and haplo.freq.min, and may yield misleading results for the rare haplotype coefficient. For this reason, we recommend that any inference made on the rare haplotypes be made with caution, if at all.

7 Methods for haplo.glm (NEW)

The latest updates to *haplo.stats* is our work to make *haplo.glm* to act similar to a glm object with methods to compare and assess model fits. In this section we describe the challenges and caveats

of defining these methods for a haplo.glm object and show how to use them.

7.1 fitted.values

A challenge when defining methods for *haplo.glm* is that we account for the ambiguity in a persons haplotype pair. To handle this in the glm framework, the response and non-haplotype covariates are expanded for each person with a posterior probability of the haplotype given their genotype as a weight. The returned object from *haplo.glm* looks somewhat like a regular glm, but the model matrix, response, and thus the fitted values, are all expanded. Users who want to work with the expanded versions of those items are welcome to access them from the returned object.

We now provide a method to get the fitted values for each person, fitted.haplo.glm. These collapsed fitted values are calculated by a weighted sum of the expanded fitted values for each person where the weights are the posterior probabilities of the person's expanded haplotype pairs.

7.2 residuals

The residuals within the *haplo.glm* object are also expanded for the haplotype pairs for subjects. We provide *residuals.haplo.glm* to get the collapsed deviance, pearson, working, and response residuals for each person. Because we have not implemented a predict method for *haplo.glm*, the method does not calculate partial residuals.

7.3 vcov

We provide vcov.haplo.glm as a method to get the variance-covariance matrix of model parameters in the haplo.glm object. Unlike the standard glm object, this matrix is computed and retained in the returned object. We do this because the model parameters are the model coefficients and the haplotype frequencies, and it is computationally-intensive to compute.

We show how to get the variance matrix for all the parameters and for only the model coefficients.

```
> varmat <- vcov(fit.gaus)</pre>
> dim(varmat)
[1] 182 182
> varmat <- vcov(fit.gaus, freq=FALSE)</pre>
> dim(varmat)
[1] 9 9
> print(varmat, digits=2)
               (Intercept)
                                male geno.glm.17 geno.glm.34 geno.glm.77
(Intercept)
                                          -0.0674
                                                       -0.0544
                                                                     -0.0526
                     0.118 -0.01513
male
                    -0.015
                             0.02409
                                           0.0065
                                                        -0.0022
                                                                     -0.0038
geno.glm.17
                    -0.067
                             0.00646
                                           0.1897
                                                        0.0335
                                                                      0.0206
geno.glm.34
                    -0.054 -0.00218
                                           0.0335
                                                        0.1179
                                                                      0.0226
geno.glm.77
                    -0.053 -0.00375
                                           0.0206
                                                        0.0226
                                                                      0.1305
```

geno.glm.78	-0.051	0.00082	0.0322	0.0275 0.0	260
geno.glm.100	-0.059	0.00674	0.0370	0.0298 0.0	204
geno.glm.138	-0.059	0.00362	0.0307	0.0254 0.0	256
geno.glm.rare	-0.058	0.00142	0.0320	0.0283 0.0	278
	geno.glm.78	geno.glm.100	geno.glm.138	geno.glm.rare	
(Intercept)	-0.05123	-0.0595	-0.0587	-0.0583	
male	0.00082	0.0067	0.0036	0.0014	
geno.glm.17	0.03217	0.0370	0.0307	0.0320	
geno.glm.34	0.02753	0.0298	0.0254	0.0283	
geno.glm.77	0.02602	0.0204	0.0256	0.0278	
geno.glm.78	0.14732	0.0285	0.0220	0.0248	
geno.glm.100	0.02853	0.1327	0.0214	0.0281	
geno.glm.138	0.02195	0.0214	0.0920	0.0288	
geno.glm.rare	0.02478	0.0281	0.0288	0.0331	

7.4 anova and Model Comparison

We use the *anova.glm* method as a framework for *anova.haplo.glm* to allow comparisons of model fits. We limit the model comparisons to multiple nested model fits, which requires that each model to be compared is either a *haplo.glm* or *glm* fitted object. We eliminate the functionality of testing sub-models of a single fit because removal of a single covariate would require re-fitting of the reduced model to get updated coefficient and haplotype frequency estimates with a maximized log-likelihood. We decided to simplify the usage and require that all models to be compared are fully fitted.

As with the *anova.glm* method, it is difficult to check for truly nested models, so we pass the responsibility on to the user. We discuss some of the requirements.

One type of two-model comparison is between models with haplotypes (expanded subjects) and a reduced model without haplotypes. We check for the same sample size in these models by comparing the collapsed sample size from a *haplo.glm* fit to the sample size from the *glm* fit, which we remind users is only a loose check of model comparability.

The other comparison of two models in *anova.haplo.glm* is to compare two models that contain the same genotypes, and inherently the same haplotypes. This is more tricky because a subject may not have the same expanded set of possible haplotype pairs across two fits of *haplo.glm* unless the same seed is set before both fits. Even if a seed is the same, the other effects that are different between the two models will affect the haplotype frequency estimates, and may still result in a different expansion of haplotype pairs per subject. Our check of the collapsed sample size for the two models still applies with the same pitfalls, but a better assurance of model comparability is to use the same seed.

In the haplo.glm fit we provide the likelihood ratio test of the null model against the full model, which is the most appropriate test available for haplo.glm objects, but it is difficult to compare the log-likelihood across two haplo.glm fits. Therefore, we remain consistent with glm model comparison [10], and use the difference in deviance to compare models. Furthermore, we restrict the asymptotic test for model comparison to be the χ^2 test for goodness of fit.

Below we show how to get the LRT from the *fit.gaus* result, then show how to compare some of the nested models fit above, including a regular glm fit of $y \sim male$. The anova method requires the nested model to be given first, and any anova with a *haplo.glm* object should explicitly call

```
anova.haplo.glm.
> fit.gaus$lrt
$lrt
[1] 22.34062
$df
[1] 8
> glmfit.gaus <- glm(y<sup>male</sup>, family="gaussian", data=glm.data)
> anova.haplo.glm(glmfit.gaus, fit.gaus)
Analysis of Deviance Table
Model 1: y ~ male
Model 2: y ~ male + geno.glm
 Resid. Df Resid. Dev Df Deviance Pr(>Chi)
1
        218
                297.00
                267.88 7 29.114 0.001752 **
        211
2
___
Signif. codes: 0 âĂŸ***âĂŹ 0.001 âĂŸ**âĂŹ 0.01 âĂŸ*âĂŹ 0.05 âĂŸ.âĂŹ 0.1 âĂŸ âĂŹ 1
> anova.haplo.glm(fit.gaus, fit.inter)
Analysis of Deviance Table
Model 1: y ~ male + geno.glm
Model 2: y ~ male * geno.glm
 Resid. Df Resid. Dev Df Deviance Pr(>Chi)
        211
                267.88
1
        204
                259.82 7 8.0631
                                     0.5017
2
> anova.haplo.glm(glmfit.gaus, fit.gaus, fit.inter)
Analysis of Deviance Table
Model 1: y ~ male
Model 2: y ~ male + geno.glm
Model 3: y ~ male * geno.glm
 Resid. Df Resid. Dev Df Deviance Pr(>Chi)
1
        218
                297.00
        211
                267.88 7 29.1137 0.001804 **
2
3
        204
               259.82 7 8.0631 0.501696
Signif. codes: 0 âĂŸ***âĂŹ 0.001 âĂŸ**âĂŹ 0.01 âĂŸ*âĂŹ 0.05 âĂŸ.âĂŹ 0.1 âĂŸ âĂŹ 1
```

8 Extended Applications

The following functions are designed to wrap the functionality of the major functions in Haplo Stats into other useful applications.

8.1 Combine Score and Group Results: haplo.score.merge

When analyzing a qualitative trait, such as binary, it can be helpful to align the results from *haplo.score* with *haplo.group*. To do so, use the function *haplo.score.merge*, as illustrated in the following example:

```
> # merge haplo.score and haplo.group results
> merge.bin <- haplo.score.merge(score.bin, group.bin)
> print(merge.bin, nlines=10)
```

Haplotype Scores, p-values, and Frequencies By Group

	DQB	DRB	В	Hap.Score	p.val	Hap.Freq	y.bin.0	y.bin.1
1	62	2	7	-2.19387	0.02824	0.05098	0.06789	0.01587
2	51	1	35	-1.58421	0.11315	0.03018	0.03754	0.00907
3	63	13	7	-1.56008	0.11874	0.01655	0.02176	NA
4	21	7	7	-1.47495	0.14023	0.01246	0.01969	NA
5	32	4	7	-1.00091	0.31687	0.01678	0.02628	0.00794
6	32	4	62	-0.67990	0.49657	0.02349	0.01911	NA
7	51	1	27	-0.66509	0.50599	0.01505	0.01855	0.00907
8	31	11	35	-0.58380	0.55936	0.01754	0.01982	0.01587
9	31	11	51	-0.43721	0.66196	0.01137	0.01321	NA
10	51	1	44	0.00826	0.99341	0.01731	0.01595	0.00000

Explanation of Results

The first column is a row index, the next columns (3 in this example) illustrate the haplotype, the Hap.Score column is the score statistic and p.val the corresponding χ^2 p-value. Hap.Freq is the haplotype frequency for the total sample, and the remaining columns are the estimated haplotype frequencies for each of the group levels (y.bin in this example). The default print method only prints results for haplotypes appearing in the haplotypes from the haplotype output. To view all haplotypes, use the print option all.haps=TRUE, which prints all haplotypes from the haplo.group output. The output is ordered by the score statistic, but the order.by parameter can specify ordering by haplotypes or by haplotype frequencyies.

8.2 Case-Control Haplotype Analysis: haplo.cc

It is possible to combine the results of haplo.score, haplo.group, and haplo.glm for case-control data, all performed within haplo.cc. The function peforms a score test and a glm on the same haplotypes. The parameters that determine which haplotypes are used are haplo.min.count and

haplo.freq.min, which are set in the control parameter, as done for haplo.glm. This is a change from previous versions, where haplo.min.count was in the parameter list for haplo.cc.

Below we run haplo.cc setting the minimum haplotype frequency at 0.02. The print results are shown, in addition to the names of the objects stored in the cc.hla result.

```
> # demo haplo.cc where haplo.min.count is specified
> # use geno, and this function prepares it for haplo.glm
> y.bin <- 1*(hla.demo$resp.cat=="low")</pre>
> cc.hla <- haplo.cc(y=y.bin, geno=geno, locus.label = label,</pre>
                control=haplo.glm.control(haplo.freq.min=.02))
+
> ##, em.c=haplo.em.control(iseed=10))) perhaps not needed??
>
> print(cc.hla, nlines=25, digits=2)
_____
                     Global Score Statistics
_____
global-stat = 29, df = 8, p-val = 0.00029
_____
                   Counts for Cases and Controls
_____
control
        case
   157
          63
_____
       Haplotype Scores, p-values, Hap-Frequencies (hf), and Odds
                        Ratios (95% CI)
_____
   DQB DRB B Hap-Score
                    p-val pool.hf control.hf case.hf glm.eff OR.lower
147
   62
       2 7
             -2.103 0.03546 0.0490
                                   0.0679 0.0159
                                                 Eff
                                                      0.0138
   51
       1 35
              -1.583 0.11344 0.0302
                                   0.0376 0.0089
                                                 Eff
98
                                                      0.0097
       4 7
                                   0.0263 0.0079
78
   32
             -1.393 0.16349 0.0227
                                                 Eff
                                                      0.0047
77
   32
       4 62
             -0.496 0.62001 0.0212
                                   0.0191
                                                 Eff
                                                      0.0517
                                            NA
76
   32
       4 60
             0.028 0.97762 0.0307
                                   0.0315 0.0238
                                                 Eff
                                                      0.0763
16
   21
       7 44
             1.069 0.28516 0.0217
                                   0.0175 0.0476
                                                 Eff
                                                      0.1253
                                   0.0150 0.0635
52
   31
       4 44
              2.516 0.01186 0.0285
                                                 Eff
                                                      0.3425
              3.776 0.00016 0.1042
11
   21
       3 8
                                   0.0693 0.1905
                                                 Base
                                                         NA
1
   21
       1 8
                 NA
                       NA 0.0023
                                   0.0033
                                            NA
                                                   R
                                                      0.1443
2
   21 10 8
                 NA
                       NA 0.0023
                                   0.0032
                                            NA
                                                   R
                                                      0.1443
3
   21
       2 18
                 NA
                       NA 0.0023
                                   0.0032
                                            NA
                                                   R
                                                      0.1443
4
   21
       2 7
                       NA 0.0023
                                   0.0032
                 NA
                                            NA
                                                   R
                                                      0.1443
5
   21
       3 18
                 NA
                       NA 0.0046
                                   0.0067
                                            NA
                                                   R
                                                      0.1443
6
   21
       3 35
                 NA
                       NA 0.0057
                                   0.0065
                                                      0.1443
                                            NA
                                                   R
7
   21
       3 44
                 NA
                       NA 0.0036
                                   0.0033 0.0159
                                                   R
                                                      0.1443
```

8 9 10 12 13 14 15 17 18 19	21 21 21 21 21 21 21 21 21 21 21	 3 49 3 57 3 70 4 62 7 13 7 18 7 35 7 45 7 50 7 57 	NA N NA N NA N NA N NA N NA N NA N	IA 0.002 IA 0.002	14 NA 13 NA 15 0.0064 18 NA 15 NA 15 NA 15 NA 15 NA 15 NA 15 NA 14 NA 13 0.0032 15 0.0032	NA NA 0.0238 NA 0.0079 NA 0.0079	R R R R R R R R R	0.1443 0.1443 0.1443 0.1443 0.1443 0.1443 0.1443 0.1443 0.1443 0.1443 0.1443
		OR.upper						
147	0.072							
98	0.086							
78	0.058							
77	0.281							
76	0.318							
16	0.661	3.48						
52	1.318	5.07						
11	1.000	NA						
1	0.290	0.58						
2	0.290	0.58						
3	0.290	0.58						
4	0.290							
5	0.290							
6	0.290							
7	0.290							
8	0.290							
9	0.290							
10	0.290							
12	0.290							
13	0.290							
14	0.290							
15	0.290							
17	0.290							
	0.290							
19	0.290	0.58						
> n	ames(c	c.hla)						
	"cc.d "excl	f" ude.subj"	"group.count	" "scor	e.lst" "f	it.lst"	"ci.p	rob"

Explanation of Results

First, from the names function we see that *cc.hla* also contains *score.lst* and *fit.lst*, which are the *haplo.score* and *haplo.glm* objects, respectively. For the printed results of *haplo.cc*, first are the

global statistics from *haplo.score*, followed by cell counts for cases and controls. The last portion of the output is a data frame containing combined results for individual haplotypes:

- Hap-Score: haplotype score statistic
- **p-val:** haplotype score statistic p-value
- sim p-val: (if simulations performed) simulated p-value for the haplotype score statistic
- **pool.hf:** haplotype frequency for the pooled sample
- control.hf: haplotype frequencies for the control sample only
- **case.hf:** haplotype frequencies for the case sample only
- glm.eff: one of three ways the haplotype appeared in the glm model: *Eff*: modeled as an effect; *Base*: part of the baseline; and *R*: a rare haplotype, included in the effect of pooled rare haplotypes
- OR.lower: Odds Ratio confidence interval lower limit
- OR: Odds Ratio for each effect in the model
- **OR.upper:** Odds Ratio confidence interval upper limit

Significance levels are indicated by the p-values for the score statistics, and the odds ratio (OR) confidence intervals for the haplotype effects. Note that the Odds Ratios are effect sizes of haplotypes, assuming haplotype effects are multiplicative. Since this last table has many columns, lines are wrapped in the output in this manual. You can align wrapped lines by the haplotype code which appears on the far left. Alternatively, instruct the print function to only print *digits* significant digits, and set the width settings for output in your session using the *options()* function.

8.3 Score Tests on Sub-Haplotypes: haplo.score.slide

To evaluate the association of sub-haplotypes (subsets of alleles from the full haplotype) with a trait, the user can evaluate a "window" of alleles by *haplo.score*, and slide this window across the entire haplotype. This procedure is implemented by the function *haplo.score.slide*. To illustrate this method, we use all 11 loci in the demo data, *hla.demo*.

First, make the geno matrix and the locus labels for the 11 loci. Then use haplo.score.slide for a window of 3 loci (n.slide=3), which will slide along the haplotype for all 9 contiguous subsets of size 3, using the previously defined gaussian trait resp.

	<pre>start.loc</pre>	<pre>score.global.p</pre>	<pre>global.p.sim</pre>	max.p.sim
1	1	0.21550	NA	NA
2	2	0.09366	NA	NA
3	3	0.39042	NA	NA
4	4	0.48771	NA	NA
5	5	0.13747	NA	NA
6	6	0.14925	NA	NA
7	7	0.11001	NA	NA
8	8	0.00996	NA	NA
9	9	0.04255	NA	NA

Explanation of Results

The first column is the row index of the nine calls to haplo.score, the second column is the number of the starting locus of the sub-haplotype, the third column is the global score statistic p-value for each call. The last two columns are the simulated p-values for the global and maximum score statistics, respectively. If you specify simulate=TRUE in the function call, the simulated p-values would be present.

8.3.1 Plot Results from haplo.score.slide

The results from haplo.score.slide can be easily viewed in a plot shown in Figure 2 below. The x-axis has tick marks for each locus, and the y-axis is the $-log_{10}(pval)$. To select which p-value to plot, use the parameter pval, with choices "global", "global.sim", and "max.sim" corresponding to p-values described above. If the simulated p-values were not computed, the default is to plot the global p-values. For each p-value, a horizontal line is drawn at the height of $-log_{10}(pval)$ across the loci over which it was calculated. For example, the p-value score.global.p = 0.009963 for loci 8-10 is plotted as a horizontal line at y = 2.002 spanning the 8^{th} , 9^{th} , and 10^{th} x-axis tick marks.



> # plot global p-values for sub-haplotypes from haplo.score.slide > plot.haplo.score.slide(score.slide.gaus)

Figure 2: Global p-values for sub-haplotypes; Gaussian Response

8.4 Scanning Haplotypes Within a Fixed-Width Window: haplo.scan

Another method to search for a candidate locus within a genome region is haplo.scan, an implementation of the method proposed in Cheng et al. 2005 [11]. This method searches for a region for which the haplotypes have the strongest association with a binary trait by sliding a window of fixed width over each marker locus, and then scans over all haplotype lengths within each window. This latter step, scanning over all possible haplotype lengths within a window, distinguishes haplo.scan from haplo.score.slide (which considers only the maximum haplotype length within a window). To account for unknown linkage phase, the function haplo.em is called prior to scanning, to create a list of haplotype pairs and posterior probabilities. To illustrate the scanning window, consider a 10-locus dataset. When placing a window of width 3 over locus 5, the possible haplotype lengths that contain locus 5 are three (loci 3-4-5, 4-5-6, and 5-6-7), two (loci 4-5 and 5-6) and one (locus 5). For each of these loci subsets a score statistic is computed, which is based on the difference between the mean vector of haplotype counts for cases and that for controls. The maximum of these score statistics, over all possible haplotype lengths within a window, is the locus-specific test statistic, or the locus scan statistic. The global test statistic is the maximum over all computed score statistics. To compute p-values, the case/control status is randomly permuted. Below we run haplo.scan on the 11-locus HLA dataset with a binary response and a window width of 3, but first we use the results of summary Geno to choose subjects with less than 50,000 haplotype pairs to speed calculations with all 11 polymorphic loci with many missing alleles.

```
> geno.11 <- hla.demo[,-c(1:4)]
> y.bin <- 1*(hla.demo$resp.cat=="low")</pre>
> hla.summary <- summaryGeno(geno.11, miss.val=c(0,NA))</pre>
> # track those subjects with too many possible haplotype pairs ( > 50,000)
> many.haps <- (1:length(y.bin))[hla.summary[,4] > 50000]
> # For speed, or even just so it will finish, make y.bin and geno.scan
> # for genotypes that don't have too many ambigous haplotypes
> geno.scan <- geno.11[-many.haps,]</pre>
> y.scan <- y.bin[-many.haps]</pre>
> # scan haplotypes for regions within width of 3 for each locus.
> # test statistic measures difference in haplotype counts in cases and controls
> # p-values are simulated for each locus and the maximum statistic,
> # we do 100 simuations here, should use default settings for analysis
>
> scan.hla <- haplo.scan(y.scan, geno.scan, width=3,</pre>
          sim.control=score.sim.control(min.sim=100, max.sim=100),
          em.control=haplo.em.control())
> print(scan.hla)
  Call:
haplo.scan(y = y.scan, geno = geno.scan,
```

```
width = 3, em.control = haplo.em.control(),
sim.control = score.sim.control(min.sim = 100,
max.sim = 100))
```

	Locus Scan-statistic Simulated P-values									
	loc-1 loc	-2 loc-3	 3 loc-4	loc-5	loc-6	loc-7	loc-8	loc-9	loc-10	loc-11
sim.p-val	0	0 () 0	0	0	0	0	0	0	0
Loci	with max	scan st	catisti	c: 2						
Max-Stat	Simulated	Global	p-value	e: 0						
	Number	of Simu	lation	s: 10	00					

Explanation of Results

In the output we report the simulated p-values for each locus test statistic. Additionally, we report the loci (or locus) which provided the maximum observed test statistic, and the Max-Stat Simulated Global p-value is the simulated p-value for that maximum statistic. We print the number of simulations, because they are performed until p-value precision criteria are met, as described in section 5.9. We would typically allow simulations to run under default parameters rather than limiting to 100 by the control parameters.

8.5 Sequential Haplotype Scan Methods: seqhap

Another approach for choosing loci for haplotype associations is by *seqhap*, as described in Yu and Schaid, 2007 [12]. The *seqhap* method performs three tests for association of a binary trait over a set of bi-allelic loci. When evaluating each locus, loci close to it are added in a sequential manner based on the Mantel-Haenszel test [13]. For each marker locus, three tests are provided:

- single locus, the traditional single-locus χ_1^2 test of association,
- sequential haplotype, based on a haplotype test for sequentially chosen loci,
- sequential sum, based on the sum of a series of conditional χ^2 statistics.

All three tests are assessed for significance with permutation p-values, in addition to the asymptotic p-value. The point-wise p-value for a statistic at a locus is the fraction of times that the statistic for the permuted data is larger than that for the observed data. The regional p-value is the chance of observing a permuted test statistic, maximized over a region, that is greater than that for the observed data.

Similar to the permutation p-values in *haplo.score* as described in section 5.9, permutations are performed until a precision threshold is reached for the regional p-values. A minimum and maximum number of permutations specified in the *sim.control* parameter list ensure a certain accuracy is met for every simulation p-value, yet having a limit to avoid infinite run-time.

Below is an example of using *seqhap* on data with case-control response for a chromosome region. First set up the binary response, y, with 0=control, 1=case, then a genotype matrix with two columns per locus, and a vector of chromosome positions. The genotype data is available in the *seqhap.dat* dataset while the chromosome positions are in *seqhap.pos*. The following example runs *seqhap* with default settings for permutations and threshold parameters.

```
> # define binary response and genotype matrix
> data(seqhap.dat)
> data(seqhap.pos)
> y <- seqhap.dat$disease</pre>
> geno <- seqhap.dat[,-1]</pre>
> # get vector with chrom position
> pos <- seqhap.pos$pos</pre>
> seqhap.out <- seqhap(y=y, geno=geno, pos=pos, miss.val=c(0,NA),
              r2.threshold=.95, mh.threshold=3.84)
+
> seqhap.out$n.sim
[1] 4973
> print(seqhap.out)
_____
                    Single-locus Chi-square Test
_____
Regional permuted P-value based on single-locus test is 0.13191
     chi.stat perm.point.p asym.point.p
loc-1 1.22062 0.27729741
                          0.26924
loc-2 1.35462 0.23245526
                          0.24447
loc-3 5.20288 0.02010859
                         0.02255
                        0.06666
0.05945
loc-4 3.36348 0.05972250
loc-5 3.55263 0.06153227
loc-6 0.39263 0.53026342
                         0.53092
loc-7 5.54913 0.01829881
                         0.01849
loc-8 3.74740 0.05469535
                         0.05289
                        0.84947
loc-9 0.03602 0.85682687
loc-10 1.99552 0.17313493
                          0.15777
_____
                        Sequential Scan
Loci Combined in Sequential Analysis
seq-loc-1 1
seq-loc-2 2 3 4 5
seq-loc-3 3 4 5
seq-loc-4 4 3
seq-loc-5 5
seq-loc-6 6 7
seq-loc-7 7
seq-loc-8 8
seq-loc-9 9
seq-loc-10 10
```

```
Sequential Haplotype Test
Regional permuted P-value based on sequential haplotype test is 0.016489
           hap.stat df perm.point.p asym.point.p
seq-loc-1
            1.22062
                    1
                        0.310878745
                                          0.26924
seq-loc-2
           24.16488 12
                        0.027950935
                                          0.01932
seq-loc-3
           19.78808
                     6
                        0.005228232
                                          0.00302
seq-loc-4
                     3
                        0.003016288
           14.95765
                                          0.00185
seq-loc-5
            3.55263
                        0.096722300
                                          0.05945
                     1
seq-loc-6
            5.45723
                     2 0.114216771
                                          0.06531
seq-loc-7
            5.54913
                     1 0.038608486
                                          0.01849
seq-loc-8
            3.74740
                     1
                        0.103961392
                                          0.05289
seq-loc-9
            0.03602
                     1
                        0.867886588
                                          0.84947
seq-loc-10
            1.99552
                        0.219384677
                                          0.15777
                     1
                               Sequential Sum Test
Regional permuted P-value based on sequential sum test is 0.0032174
           sum.stat df perm.point.p asym.point.p
seq-loc-1
            1.22062
                     1 0.3108787452
                                          0.26924
seq-loc-2
                     4 0.0008043435
                                          0.00030
           21.15360
seq-loc-3
           18.65769
                     3 0.0008043435
                                          0.00032
seq-loc-4
           14.61897
                     2 0.0020108586
                                          0.00067
seq-loc-5
            3.55263
                     1 0.1033581339
                                          0.05945
seq-loc-6
            5.43826
                     2 0.1150211140
                                          0.06593
seq-loc-7
            5.54913
                     1 0.0386084858
                                          0.01849
seq-loc-8
            3.74740
                     1 0.1041624774
                                          0.05289
seq-loc-9
            0.03602
                     1 0.8678865876
                                          0.84947
seq-loc-10
                     1 0.2193846773
            1.99552
                                          0.15777
```

Explanation of Results

The output from this example first shows *n.sim*, the number of permutations needed for precision on the regional p-values. Next, in the printed results, the first section (Single-locus Chi-square Test) shows a table with columns for single-locus tests. The table includes test statistics, permuted pvalues, and asymptotic p-values based on a χ_1^2 distribution. The second section (Sequential Scan) shows which loci are combined for association. In this example, the table shows the first locus is not combined with other loci, whereas the second locus is combined with loci 3, 4, and 5. The third section (Sequential Haplotype Test), shows the test statistics for the sequential haplotype method with degrees of freedom and permuted and asymptotic p-values. The fourth section (Sequential Sum Test) shows similar information for the sequential sum tests.

8.5.1 Plot Results from seqhap

The results from *seqhap* can be viewed in a useful plot shown in Figure 3. The plot is similar to the plot for *haplo.score.slide* results, with the x-axis having tick marks for all loci and the y-axis is the -log10() of p-value for the tests performed. For the sequential result for each locus, a horizontal line at the height of -log10(p-value) is drawn across the loci combined. The start locus is indicated by a filled triangle and other loci combined with the start locus are indicated by an asterisk or circle. The choices for pval include "hap" (sequential haplotype asymptotic p-value), "hap.sim" (sequential haplotype simulated p-value), "sum" (sequential sum asymptotic p-value), and "sum.sim" (sequential sum simulated p-value). The other parameter option is single, indicating whether to plot a line for the single-locus tests.



> # plot global p-values for sub-haplotypes from haplo.score.slide
> plot(seqhap.out, pval="hap", single=TRUE, las=2)

Figure 3: Plot p-values for sequential haplotype scan and single-locus tests

8.6 Creating Haplotype Effect Columns: haplo.design

In some instances, the desired model for haplotype effects is not possible with the methods given in *haplo.glm*. Examples include modeling just one haplotype effect, or modeling an interaction of haplotypes from different chromosomes, or analyzing censored data. To circumvent these limitations, we provide a function called *haplo.design*, which will set up an expected haplotype design matrix from a *haplo.em* object, to create columns that can be used to model haplotype effects in other modeling functions.

The function haplo.design first creates a design marix for all pairs of haplotypes over all subjects, and then uses the posterior probabilities to create a weighted average contribution for each subject, so that the number of rows of the final design matrix is equal to the number of subjects. This is sometimes called the expectation-substitution method, as proposed by Zaykin et al. 2002 [4], and using this haplotype design matrix in a regression model is asymptotically equivalent to the score statistics from haplo.score (Xie and Stram 2005 [14]). Although this provides much flexibility, by using the design matrix in any type of regression model, the estimated regression parameters can be biased toward zero (see Lin and Zeng, 2006 [15] for concerns about the expectation-substitution method).

In the first example below, using default parameters, the returned data.frame contains a column for each haplotype that meets a minimum count in the sample *min.count*. The columns are named by the code they are assigned in *haplo.em*.

```
> # create a matrix of haplotype effect columns from haplo.em result
> hap.effect.frame <- haplo.design(save.em)
> names(hap.effect.frame)
[1] "hap.4" "hap.13" "hap.17" "hap.34" "hap.50" "hap.55" "hap
```

[1] "hap.4" "hap.13" "hap.17" "hap.34" "hap.50" "hap.55" "hap.69"
[8] "hap.77" "hap.78" "hap.99" "hap.100" "hap.102" "hap.138" "hap.140"
[15] "hap.143" "hap.155" "hap.162" "hap.165"

> hap.effect.frame[1:10,1:8]

	hap.4	hap.13	hap.17	hap.34	hap.50	hap.55	hap.69	hap.77
1	0	0.000000	0.000000	0	0	0	0	0
2	0	0.1253234	0.8746766	0	0	0	0	0
3	0	0.000000	0.000000	0	0	0	0	0
4	0	0.2862131	0.7137869	0	0	0	0	0
5	0	0.000000	0.000000	0	0	1	0	0
6	1	0.000000	1.000000	0	0	0	0	0
7	0	0.000000	0.000000	0	0	0	0	0
8	0	0.000000	0.000000	0	0	0	0	0
9	0	0.000000	0.000000	0	0	0	0	0
10	0	0.000000	0.000000	0	0	0	0	0

Additionally, *haplo.design* gives the user flexibility to make a more specific design matrix with the following parameters:

• hapcodes: codes assigned in the haplo.em object, the only haplotypes to be made into effects

- haplo.effect: the coding of haplotypes as additive, dominant, or recessive
- haplo.base: code for the baseline haplotype
- min.count: minimum haplotype count

This second example below creates columns for specific haplotype codes that were most interesting in *score.gaus.add*, haplotypes with alleles 21-3-8 and 62-2-7, corresponding to codes 4 and 138 in *haplo.em*, respectively. Assume we want to test their individual effects when they are coded with *haplo.effect="dominant"*.

```
> # create haplotype effect cols for haps 4 and 138
> hap4.hap138.frame <- haplo.design(save.em, hapcodes=c(4,138),</pre>
                                   haplo.effect="dominant")
+
> hap4.hap138.frame[1:10,]
  hap.4
           hap.138
       0 0.000000
1
2
       0 0.8746766
3
       0 0.000000
4
       0 0.0000000
5
       0 0.000000
6
       1 0.000000
7
       0 1.0000000
8
       0 0.0000000
9
       0 0.1358696
10
       0 0.0000000
> dat.glm <- data.frame(resp, male, age,</pre>
+
                        hap.4=hap4.hap138.frame$hap.4,
                        hap.138=hap4.hap138.frame$hap.138)
> glm.hap4.hap138 <- glm(resp ~ male + age + hap.4 + hap.138,
                         family="gaussian", data=dat.glm)
+
> summary(glm.hap4.hap138)
Call:
glm(formula = resp ~ male + age + hap.4 + hap.138, family = "gaussian",
    data = dat.glm)
Deviance Residuals:
    Min
                1Q
                      Median
                                    ЗQ
                                              Max
-2.32614 -1.07489 -0.06559
                              1.04483
                                          2.39044
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
                        0.229577
                                   8.336 9.11e-15 ***
(Intercept) 1.913834
male
             0.048588
                       0.155290
                                   0.313
                                          0.7547
```

age -0.002651 0.011695 -0.227 0.8209 hap.4 -0.405530 0.195857 -2.071 0.0396 * hap.138 0.584480 0.261763 2.233 0.0266 * ---Signif. codes: 0 âĂŸ***âĂŹ 0.001 âĂŸ**âĂŹ 0.01 âĂŸ*âĂŹ 0.05 âĂŸ.âĂŹ 0.1 âĂŸ âĂŹ 1

(Dispersion parameter for gaussian family taken to be 1.318277)

Null deviance: 297.01 on 219 degrees of freedom Residual deviance: 283.43 on 215 degrees of freedom AIC: 692.07

Number of Fisher Scoring iterations: 2

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Appendix

A Counting Haplotype Pairs When Marker Phenotypes Have Missing Alleles

The following describes the process for counting the number of haplotype pairs that are consistent with a subject's observed marker phenotypes, allowing for some loci with missing data. Note that we refer to marker phenotypes, but our algorithm is oriented towards typical markers that have a one-to-one correspondence with their genotypes. We first describe how to count when none of the loci have missing alleles, and then generalize to allow loci to have either one or two missing alleles. When there are no missing alleles, note that homozygous loci are not ambiguous with respect to the underlying haplotypes, because at these loci the underlying haplotypes will not differ if we interchange alleles between haplotypes. In contrast, heterozygous loci are ambiguous, because we do not know the haplotype origin of the distinguishable alleles (i.e., unknown linkage phase). However, if there is only one heterozygous locus, then it doesn't matter if we interchange alleles, because the pair of haplotypes will be the same. In this situation, if parental origin of alleles were known, then interchanging alleles would switch parental origin of haplotypes, but not the composition of the haplotypes. Hence, ambiguity arises only when there are at least two heterozygous loci. For each heterozygous locus beyond the first one, the number of possible haplotypes increases by a factor of 2, because we interchange the two alleles at each heterozygous locus to create all possible pairs of haplotypes. Hence, the number of possible haplotype pairs can be expressed as 2^x , where x = H - 1, if H (the number of heterozygous loci) is at least 2, otherwise x = 0.

Now consider a locus with missing alleles. The possible alleles at a given locus are considered to be those that are actually observed in the data. Let a_i denote the number of distinguishable alleles at the locus. To count the number of underlying haplotypes that are consistent with the observed and missing marker data, we need to enumerate all possible genotypes for the loci with missing data, and consider whether the imputed genotypes are heterozygous or homozygous.

To develop our method, first consider how to count the number of genotypes at a locus, say the i^{th} locus, when either one or two alleles are missing. This locus could have either a homozygous or heterozygous genotype, and both possibilities must be considered for our counting method. If the locus is considered as homozygous, and there is one allele missing, then there is only one possible genotype; if there are two alleles missing, then there are a_i possible genotypes. A function to perform this counting for homozygous loci is denoted $f(a_i)$. If the locus is considered as heterozygous, and there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes; if there are two alleles missing, then there are $a_i - 1$ possible genotypes. A function to perform this counting for homozygous loci is denoted $g(a_i)$. These functions and counts are summarized in Table A.1.

Table A.1: Factors for when a locus having missing allele(s) is counted as homozygous(f()) or heterozygous(g())

Now, to use these genotype counting functions to determine the number of possible haplotype pairs, first consider a simple case where only one locus, say the i^{th} locus, has two missing alleles. Suppose that the phenotype has H heterozygous loci (H is the count of heterozygous loci among those without missing data). We consider whether the locus with missing data is either homozygous or heterozygous, to give the count of possible haplotype pairs as

Numbe	er of	Homozygous	Heterozygous
missing all	leles	function $f(a_i)$	function $g(a_i)$
	1	1	$a_i - 1$
	2	a_i	$\frac{a_i(a_i-1)}{2}$

$$a_i 2^x + \left[\frac{a_i(a_i-1)}{2}\right] 2^{x+1} \tag{1}$$

where again x = H - 1 if H is at least 2, otherwise x = 0. This special case can be represented by our more general genotype counting functions as

$$f(a_i) 2^x + g(a_i) 2^{x+1}$$
(2)

When multiple loci have missing data, we need to sum over all possible combinations of heterozygous and homozygous genotypes for the incomplete loci. The rows of Table A.2 below present these combinations for up to m = 3 loci with missing data. Note that as the number of heterozygous loci increases (across the columns of Table A.2), so too does the exponent of 2. To calculate the total number of pairs of haplotypes, given observed and possibly missing genotypes, we need to sum the terms in Table A.2 across the appropriate row. For example, with m = 3, there are eight terms to sum over. The general formulation for this counting method can be expressed as

$$TotalPairs = \sum_{j=0}^{m} \sum_{combo} C(combo, j)$$
(3)

where combo is a particular pattern of heterozygous and homozygous loci among the loci with missing values (e.g., for m = 3, one combination is the first locus heterozygous and the 2^{nd} and 3^{rd} third as homozygous), and C(combo, j) is the corresponding count for this pattern when there are *i* loci that are heterozygous (e.g., for m = 3 and j = 1, as illustrated in Table A.2). **Table A.2:** Genotype counting terms when *m* loci have missing

alleles, grouped by number of heterozygous loci (out of m)

m	j = 0 of m	j = 1 of m	j = 2 of m	j = 3 of m
0	2^x			
1	$f(a_1)2^x$	$g(a_1)2^{x+1}$		
2	$f(a_1)f(a_2)2^x$	$g(a_1)f(a_2)2^{x+1}$	$g(a_1)g(a_2)2^{x+1}$	
		$f(a_1)g(a_2)2^{x+1}$		
3	$f(a_1)f(a_2)f(a_3)2^x$	$g(a_1)f(a_2)f(a_3)2^{x+1}$	$g(a_1)g(a_2)f(a_3)2^{x+2}$	$g(a_1)g(a_2)g(a_3)2^{x+2}$
		$f(a_1)g(a_2)f(a_3)2^{x+1}$	$g(a_1)f(a_2)g(a_3)2^{x+2}$	
		$f(a_1)f(a_2)g(a_3)2^{x+1}$	$f(a_1)g(a_2)g(a_3)2^{x+2}$	

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